This is the peer reviewed version of the following article: Martínez, S., Rueda, M. D. M., \& Illescas, M. D. (2022). Reduction of optimal calibration dimension with a new optimal auxiliary vector for calibrated estimators of the distribution function. Mathematical Methods in the Applied Sciences, 45(17), 10959-10981, which has been published in final form at https://doi.org/10.1002/mma. 8431 . This article may be used for non-commercial purposes in accordance with Wiley Terms and Conditions for Use of Self-Archived Versions. This article may not be enhanced, enriched or otherwise transformed into a derivative work, without express permission from Wiley or by statutory rights under applicable legislation. Copyright notices must not be removed, obscured or modified. The article must be linked to Wiley's version of record on Wiley Online Library and any embedding, framing or otherwise making available the article or pages thereof by third parties from platforms, services and websites other than Wiley Online Library must be prohibited.

## ARTICLE TYPE

# Reduction of optimal calibration dimension with a new optimal auxiliary vector for calibrated estimators of the distribution function. 

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#### Abstract

Summary The calibration method ${ }^{11}$ has been widely used to incorporate auxiliary information in the estimation of various parameters. Specifically, ${ }^{[2]}$ adapted this method to estimate the distribution function, although their proposal is computationally simple, its efficiency depends on the selection of an auxiliary vector of points. This work deals with the problem of selecting the calibration auxiliary vector that minimize the asymptotic variance of the calibration estimator of distribution function. The optimal dimension of the optimal auxiliary vector is reduced considerably with respect to previous studies ${ }^{[3]}$ so that with a smaller set of points the minimum of the asymptotic variance can be reached, which in turn allows to improve the efficiency of the estimates.


## KEYWORDS:

Survey sampling, distribution function, auxiliary information, calibration

## 1 | INTRODUCTION

In sample surveys, auxiliary population information is sometimes used in the estimation stage to increase the precision of the estimators of a mean or total population. Previous literature has investigated the use of auxiliary information to improve the estimation of a finite population mean, however, previous studies have considered to a lesser extent the development of efficient methods to estimate the distribution function and the finite population quantiles by incorporating the auxiliary information. The estimation of finite population distribution function is an important issue because the distribution function can be more useful than means and totals ${ }^{[4]}$.Through the finite population distribution function, parameters such as population quantiles can be obtained. More specifically, in economics, many indicators used in the poverty analisys are based on quantiles, since they analyze variables with skewed distributions such as income, and in such cases the median is a more suitable location measure than the mean. Moreover, poverty studies incorporate the analysis of wage inequality and income distribution thorugh percentile ratios ${ }^{5 / 677}$.

In the last decade, the well-known calibration estimation method to estimate the population total ${ }^{[1]}$ has been employed to develop new estimators which incorporates the auxiliary information available and it has become an important field of research in survey sampling ${ }^{[2|8| 9|10| 11}$.

Previous works ${ }^{2[12 \mid 13]}$ use different implementations of the calibration approach to obtain estimators of the distribution function and the quantiles. Under a general superpopulation model ${ }^{144}$ propose a model-calibrated estimators that is optimal under a chosen
model with respect to the anticipated variance. Although ${ }^{[14}$ considers a general sampling design, its proposal does not produce an estimator with the properties of a genuine distribution function unless the weight system is obtained by using a point $t_{0}$ for any $t$ value, which restricts the efficiency of the estimator to a neighborhood of $t_{0}$. Additionally, the proposal ${ }^{[14]}$ requires the estimation of certain superpopulation parameters that depend on the study variable, which may restrict its applicability in some cases and also require additional conditions on the sampling design to maintain the asymptotic behavior of the proposed estimator ${ }^{[15]}$.

Nonparametric regression $\left({ }^{[16} \mathrm{and}^{[2]}\right)$, is also used for model-calibration estimation of the distribution function ${ }^{17]}$ propose a new estimator for the distribution function that integrates ideas from model calibration and penalized calibration. The method ${ }^{2]}$ is computationally simple and it employs the calibration method by minimizing the chi-square distance subject to calibration equations that require the use of arbitrarily fixed values. One drawback of these estimators is that their efficiency depends on selected points. Under simple random sampling, the problem of optimal selection points in order to obtain the best estimation has been treated in previous works ${ }^{3 / 18 / 19 \mid 20}$. In fact, the work ${ }^{3 /}$ obtained the optimal dimension and the optimal auxiliary vector for the estimator of the distribution function proposed in the work ${ }^{[2]}$ and although this proposal do not generate a unique weight system that is optimal for each point $t$, it produces an estimator that is computationally simple and is a genuine distribution function that can be used directly in the estimation of quantiles and poverty measures ${ }^{21}$.

In many situations, the optimal auxiliary vector has a very high dimension, which makes the calibration process difficult and can also affect the efficiency of the estimator. Performing calibration with a high dimensional auxiliary dataset can be several problems: the variance of the calibration estimator can be increases and the optimisation procedure may fail. ${ }^{[22]}$ showed that if too many auxiliary variables are used, the bias of the calibrated estimator increases and can become nonnegligible compared to the variance (over-calibration). Recently ${ }^{23}$ theoretically prove that over-calibration may deteriorate the efficiency of the estimates. Various procedures have been suggested for variable selection. ${ }^{22}$ computed the mean squared error (MSE) for all possible subsets of quantitative auxiliary variables and then chose the one producing the smallest MSE. Later, ${ }^{244}$ used forward and stepwise selection based on the difference between the MSE of the prediction for two nested sets of variables. Alternatively, the least absolute shrinkage and selection operator (LASSO) $\left({ }^{[25}\right)$ might be considered for selecting the best subsets. Once the best set of regressors has been selected, the calibration is performed on these variables alone. Another approach to consider is that of penalised calibration $\left({ }^{[26}\right)$, which takes account of auxiliary information by attaching more or less importance according to its presumed explanatory power for the variable of interest. In a different way, ${ }^{[27}$ and ${ }^{[28]}$ suggested applying principal component analysis for quantitative auxiliary variables in order to achieve a strong dimension reduction. These works are oriented to the estimation of linear parameters.

In this work, we intend to analyze whether it is possible to reduce the optimal dimension of the auxiliary vector proposed in the previous work ${ }^{3}$. The remainder of the article is organized as follow. After introducing the problem of distribution function estimation in Section 2 with the method proposed in research work ${ }^{[2]}$ and the optimal auxiliary vector proposed in the previous work ${ }^{[3]}$, in Section 3 we will analyze the conditions under which we can reduce the dimension of the optimal auxiliary vector. Then, Section 4 proposes a new calibration estimator based on the results of Section 3. Section 5 reports the results of an extensive simulation study run on a set of synthetic and real finite populations in which the performance of the proposed class of estimators is investigated for finite size samples. Section 6 provides some conclusions.

## 2 | CALIBRATION ESTIMATION OF THE DISTRIBUTION FUNCTION AND OPTIMAL AUXILIARY VECTOR

Let $U=\{1, \ldots, N\}$ a finite population composed of $N$ different units and let $s=\{1,2, \ldots, n\}$ a random sample of size $n$ selected using a specified sampling design $p(\cdot)$ with first and second-order inclusion probabilities $\pi_{k}>0$ and $\pi_{k l}>0 k, l \in U$ respectively and $d_{k}=\pi_{k}^{-1}$ denotes the sampling design-basic weight for unit $k \in U$. Let $y_{k}$ be the study variable and $\mathbf{x}_{k}^{\prime}=\left(x_{1 k}, \ldots, x_{J k}\right)$ be a vector of auxiliary variables at unit $k$. We assume that value $\mathbf{x}_{k}$ is available for all population units whereas the value $y_{k}$ is available only for sample units. The distribution function $F_{y}(t)$ for the study variable $y$ is defined as follow:

$$
\begin{equation*}
F_{y}(t)=\frac{1}{N} \sum_{k \in U} \Delta\left(t-y_{k}\right) \tag{1}
\end{equation*}
$$

with

$$
\Delta\left(t-y_{k}\right)=\left\{\begin{array}{l}
1 \text { si } t \geq y_{k} \\
0 \text { si } t<y_{k}
\end{array}\right.
$$

A design-based estimator of the distribution function $F_{y}(t)$ is the Horvitz-Thompson estimator, defined by

$$
\begin{equation*}
\widehat{F}_{Y H T}(t)=\frac{1}{N} \sum_{k \in s} d_{k} \Delta\left(t-y_{k}\right) \tag{2}
\end{equation*}
$$

The estimator $\widehat{F}_{Y H T}(t)$ is unbiased, but it does not incorporate the auxiliary information provided by the auxiliary vector $\mathbf{x}$.
Several authors ${ }^{2|13| 29|30|}$ have incorporated the auxiliary information to obtain new estimators of $F_{y}(t)$ through the calibration method $^{[1]}$. The proposal ${ }^{[2]}$ applies the calibration procedure from a pseudo-variable

$$
\begin{gather*}
g_{k}=\left(\widehat{\beta}^{\prime} \mathbf{x}_{k} \text { for } k=1,2, \ldots N\right.  \tag{3}\\
\widehat{\beta}=\left(\sum_{k \in s} d_{k} \mathbf{x}_{k} \mathbf{x}_{k}^{\prime}\right)^{-1} \cdot \sum_{k \in s} d_{k} \mathbf{x}_{k} y_{k} \tag{4}
\end{gather*}
$$

With the variable $g$, the basic weights $d_{k}$ are replaced by new calibrated weights $\omega_{k}$ through the minimization of the chi-square distance measure

$$
\begin{equation*}
\Phi_{s}=\sum_{k \in s} \frac{\left(\omega_{k}-d_{k}\right)^{2}}{d_{k} q_{k}} \tag{5}
\end{equation*}
$$

subject to the calibration constrains

$$
\begin{equation*}
\frac{1}{N} \sum_{k \in s} \omega_{k} \Delta\left(t_{j}-g_{k}\right)=F_{g}\left(t_{j}\right) \quad j=1,2, \ldots, P \tag{6}
\end{equation*}
$$

where $F_{g}\left(t_{j}\right)$ denotes the finite distribution function of the pseudo-variable $g_{k}$ evaluated at the points $t_{j}, \quad j=1,2, \ldots, P$. We assume, with no loss in generality, $t_{1}<t_{2}<\ldots t_{P}$. The values $q_{k}$ are known positive constants unrelated to $d_{k}$.
Following ${ }^{22}$, we assume that the matrix $T$ given by:

$$
T=\sum_{k \in s} d_{k} q_{k} \Delta\left(\mathbf{t}_{\mathbf{g}}-g_{k}\right) \Delta\left(\mathbf{t}_{\mathbf{g}}-g_{k}\right)^{\prime}
$$

is nonsingular. With this calibration procedure, the calibration estimator obtained is:

$$
\begin{equation*}
\widehat{F}_{y c}(t)=\widehat{F}_{Y H T}(t)+\left(F_{g}\left(\mathbf{t}_{\mathbf{g}}\right)-\widehat{F}_{G H T}\left(\mathbf{t}_{\mathbf{g}}\right)\right)^{\prime} \cdot \widehat{D}\left(\mathbf{t}_{\mathbf{g}}\right) \tag{7}
\end{equation*}
$$

where

$$
\widehat{D}\left(\mathbf{t}_{\mathbf{g}}\right)=T^{-1} \cdot \sum_{k \in s} d_{k} q_{k} \Delta\left(\mathbf{t}_{\mathbf{g}}-g_{k}\right) \Delta\left(t-y_{k}\right)
$$

and $\widehat{F}_{G H T}\left(\mathbf{t}_{\mathbf{g}}\right)$ is the Horvitz-Thompson estimator of $F_{g}\left(\mathbf{t}_{\mathbf{g}}\right)$ evaluated at $\mathbf{t}_{\mathbf{g}}=\left(t_{1}, \ldots, t_{P}\right)^{\prime}$.
The calibration estimator $\widehat{F}_{y c}(t)$ has the following asymptotic variance ${ }^{2 /}$ :

$$
\begin{equation*}
A V\left(\widehat{F}_{y c}(t)\right)=\frac{1}{N^{2}} \sum_{k \in U} \sum_{l \in U} \Delta_{k l}\left(d_{k} E_{k}\right)\left(d_{l} E_{l}\right) \tag{8}
\end{equation*}
$$

where $E_{k}=\Delta\left(t-y_{k}\right)-\Delta\left(\mathbf{t}_{\mathbf{g}}-g_{k}\right) \cdot D\left(\mathbf{t}_{\mathbf{g}}\right)$, with

$$
\begin{equation*}
D\left(\mathbf{t}_{\mathbf{g}}\right)=\left(\sum_{k \in U} q_{k} \Delta\left(\mathbf{t}_{\mathbf{g}}-g_{k}\right) \Delta\left(\mathbf{t}_{\mathbf{g}}-g_{k}\right)^{\prime}\right)^{-1} \cdot\left(\sum_{k \in U} q_{k} \Delta\left(\mathbf{t}_{\mathbf{g}}-g_{k}\right) \Delta\left(t-y_{k}\right)\right) \tag{9}
\end{equation*}
$$

As a consequence, the behavior of the estimator $\widehat{F}_{y c}(t)$ and its precision depends on the selection of the vector $\mathbf{t}_{\mathbf{g}}$.
Previous works ${ }^{3 / 1920}$ treated, under simple random sampling without replacement and $q_{k}=c$ for all $k \in U$, the optimal selection of the vector $\mathbf{t}_{\mathbf{g}}$ in order to minimize the asymptotic variance (8). In fact, ${ }^{3}$ established the optimal dimension of $\mathbf{t}_{\mathbf{g}}$ and its optimal value, for a given value $t$, through the definition of the sets:

$$
\begin{equation*}
A_{t}=\left\{g_{k}: k \in U ; y_{k} \leq t\right\}=\left\{a_{1}^{t}, a_{2}^{t}, \ldots, a_{M_{t}}^{t}\right\} \quad \text { with } \quad a_{h}^{t}<a_{h+1}^{t} \quad \text { for } \quad h=1, \ldots, M_{t}-1 \tag{10}
\end{equation*}
$$

where $M_{t}$ is the number of elements in the set $A_{t}$ and

$$
\begin{equation*}
B_{t}=\left\{b_{1}^{t}, b_{2}^{t}, \ldots, b_{M_{t}}^{t}\right\} \tag{11}
\end{equation*}
$$

with

$$
\begin{aligned}
& b_{1}^{t}=\max _{l \in U_{1}}\left\{g_{l}\right\} \quad \text { where } U_{1}=\left\{l \in U: g_{l}<a_{1}^{t}\right\} \\
& b_{h}^{t}=\max _{l \in U_{h}}\left\{g_{l}\right\} \quad \text { where } U_{h}=\left\{l \in U: a_{h-1}^{t}<g_{l}<a_{h}^{t}\right\} \quad h=2,3, \ldots, M_{t}
\end{aligned}
$$

and $b_{h}^{t}<b_{h+1}^{t}$ for $h=1, \ldots, M_{t}-1$.
Thus, ${ }^{3}$ established that the auxiliary vector $\mathbf{t}_{g}$ has optimal dimension $P=2 M_{t}$ if $b_{h}^{t}$ exists for $h=1, \ldots, M_{t}$ and the optimal value of $\mathbf{t}_{g}$ is given by

$$
\begin{equation*}
\mathbf{t}_{\mathbf{O P T}}(t)=\left(b_{1}^{t}, a_{1}^{t}, \ldots, b_{M_{t}}^{t}, a_{M_{t}}^{t}\right) \tag{12}
\end{equation*}
$$

If there are some values $j_{1}^{t}, j_{2}^{t}, \ldots j_{p_{t}}^{t} \in\left\{1, \ldots, M_{t}\right\}$; such as $b_{j_{h}}^{t}$ does not exits for $h=1,2, \ldots p_{t}$ with $p_{t} \leq M_{t}$ and $j_{h}^{t} \neq j_{q}^{t}$ if $h \neq q$, the optimal dimension is given by $P=2 M_{t}-p_{t}$ and the optimal auxiliary vector $\mathbf{t}_{\mathbf{O P}}$ is:

$$
\begin{equation*}
\mathbf{t}_{\mathbf{O P}}(t)=\left(b_{1}^{t}, a_{1}^{t} \ldots, b_{j_{1}-1}^{t}, a_{j_{1}-1}^{t}, a_{j_{1}}^{t}, b_{j_{1}+1}^{t}, \ldots, b_{j_{h}-1}^{t}, a_{j_{h}-1}^{t}, a_{j_{h}}^{t}, b_{j_{h}+1}^{t}, \ldots b_{M_{t}}^{t}, a_{M_{t}}^{t}\right) \tag{13}
\end{equation*}
$$

In the next section, we will analyze if the minimum of the asymptotic variance can be reached with a vector of less dimension and we will establish conditions under which the dimension of the optimal vector $\mathbf{t}_{\mathbf{O P T}}(t)$, can be reduced under simple random sampling without replacement.

## 3 | DIMENSION REDUCTION OF THE OPTIMAL AUXILIARY VECTOR

In this section, we will analyze the conditions under which the dimension of the optimal vector $\mathbf{t}_{\mathbf{O P T}}(t)$ can be reduced, that is, we will analyze the existence of a vector with a smaller dimension than $\mathbf{t}_{\mathbf{O P T}}(t)$ that allows obtaining the minimum value of the asymptotic variance of the estimator $\widehat{F}_{y c}(t)$.

For the minimization of the asymptotic variance (8), we consider it as a function of a vector $\gamma=\left(\gamma_{1}, \ldots, \gamma_{P}\right)$ of dimension $P$ :

$$
\begin{equation*}
A V\left(\widehat{F}_{y c}(t)\right)=\frac{1}{N^{2}} \sum_{k \in U} \sum_{l \in U} \Delta_{k l}\left(d_{k} \Gamma_{k}\right)\left(d_{l} \Gamma_{l}\right) \tag{14}
\end{equation*}
$$

with $\Gamma_{k}=\Delta\left(t-y_{k}\right)-\Delta\left(\boldsymbol{\gamma}-g_{k}\right) \cdot D(\boldsymbol{\gamma})$, with $D(\boldsymbol{\gamma})$ given by 9 .

Following ${ }^{[20}$, under simple random sampling without replacement and $q_{k}=c$ for all units in the population, the minimization of (14) is equivalent to the minimization of the function:

$$
\begin{equation*}
Q_{t}(\gamma)=Q_{t}\left(\gamma_{1}, \ldots, \gamma_{P}\right)=2 N F_{y}(t) \cdot K_{t}\left(\gamma_{P}\right)-\sum_{j=1}^{P} \frac{\left(K_{t}\left(\gamma_{j}\right)-K_{t}\left(\gamma_{j-1}\right)\right)^{2}}{\left(F_{g}\left(\gamma_{j}\right)-F_{g}\left(\gamma_{j-1}\right)\right)}-\left(K_{t}\left(\gamma_{P}\right)\right)^{2} \tag{15}
\end{equation*}
$$

with $K_{t}\left(\gamma_{j}\right)=\sum_{k \in U} \Delta\left(\gamma_{j}-g_{k}\right) \Delta\left(t-y_{k}\right)$, where we suppose that $F_{g}\left(\gamma_{0}\right)=0$ and $K_{t}\left(\gamma_{0}\right)=0$.
As mentioned above, ${ }^{[3]}$ established the optimal dimension $P$ and the minimum of $\sqrt{14}$ is reached at $\boldsymbol{\gamma}=\mathbf{t}_{\mathbf{o p}}(t)$. However, there are cases where the optimal dimension has a high value so the calibration procedure has a plenty of constraints which raises the computational cost for calculating the estimator. For example, if we consider $t=y_{\max }$ where

$$
y_{\max }=\max _{k \in U} y_{k}
$$

the optimal auxiliary vector $\mathbf{t}_{\mathbf{O P}}(t)=\left(a_{1}, a_{2}, \ldots, a_{M}\right)$ can be reduced to the auxiliary vector $\boldsymbol{\gamma}=\left(a_{M}\right)$ (see Appendix A.1). Consequently, the optimal dimension can be reducted from $M$ to 1 .

In a similar way, we try to reduce the dimension of the auxiliary vector to reach the minimum of $Q_{t}(\boldsymbol{\gamma})$. For it, given a value $t$ for which we want to estimate $F_{y}(t)$, we consider the sets $A_{M}, A_{t}$ and $B_{t}$ given by $\left.\sqrt{\text { A1 }} ; 10\right]$ and 11$)$ respectively and for each $a_{i} \in A_{M}$ we define:

$$
r_{i}=\text { Frequency of the } a_{i}
$$

For the value $t$, we have:

$$
A_{t}=\left\{a_{1}^{t}, a_{2}^{t}, \ldots, a_{M_{t}}^{t}\right\}=\left\{a_{f_{1}^{t}}, a_{f_{2}^{t}}, \ldots, a_{f_{M_{t}}^{t}}\right\}
$$

where

$$
\left\{f_{1}^{t}, f_{2}^{t}, \ldots, f_{M_{t}}^{t}\right\} \subseteq\{1,2, \ldots, M\} \text { and } f_{1}^{t}<f_{2}^{t}<\cdots<f_{M_{t}}^{t}
$$

Similarly, we consider the following set:

$$
C_{t}=\left\{g_{k}: k \in U ; y_{k}>t\right\}=\left\{c_{1}^{t}, c_{2}^{t}, \ldots, c_{S_{t}}^{t}\right\}=\left\{a_{l_{1}^{t}}, a_{l_{2}^{t}}, \ldots, a_{l_{s_{t}}}\right\}
$$

with

$$
\left\{l_{1}^{t}, l_{2}^{t}, \ldots, l_{S_{t}}^{t}\right\} \subseteq\{1,2, \ldots, M\} \text { and } l_{1}^{t}<l_{2}^{t}<\cdots<l_{S_{t}}^{t}
$$

It is clear that $A_{t} \cup C_{t}=A_{M}$ and since for two different units $k$ and $j$ can be possible that $g_{j}=g_{k}=a_{i}$ and $y_{k}>t$ and $y_{j}<t$, not necessarily $A_{t} \cap C_{t}=\emptyset$. For the sets $A_{t}$ and $C_{t}$ we define:

$$
\begin{aligned}
p_{i}^{t} & =\text { Frequency of the } a_{i}^{t} \text { in } A_{t} \\
q_{i}^{t} & =\text { Frequency of the } c_{i}^{t} \text { in } C_{t} .
\end{aligned}
$$

Next, we consider the following sets:

$$
\begin{gather*}
D_{t}=\left\{c_{i} \in C_{t}: q_{i}^{t}=r_{i}\right\}  \tag{16}\\
Z_{t}=\left\{a_{i}^{t} \in A_{t}: q_{i}^{t}=0\right\}=\left\{a_{i}^{t} \in A_{t}: a_{i}^{t} \notin C_{t}\right\}=A_{t}-C_{t}  \tag{17}\\
F_{t}=\left\{a_{i}^{t} \in A_{t}: 0<q_{i}^{t}<r_{i}\right\} . \tag{18}
\end{gather*}
$$

It is easy to see that $D_{t}=A_{M}-A_{t}$ and consequently $A_{t} \cap D_{t}=\emptyset$. Futhermore, $B_{t} \subseteq D_{t} ; A_{t}=Z_{t} \cup F_{t}$ and $Z_{t} \cap F_{t}=\emptyset$.
Firstly, if we suppose that $D_{t}=A_{M}$, we have $A_{t}=\emptyset$ and consequently $y_{k}>t, \forall k \in U$. In this case $F_{y}(t)=0$ and we can calibrate with any auxiliary vector since

$$
\widehat{F}_{y c}(t)=\frac{1}{N} \sum_{k \in s} \omega_{k} \Delta\left(t-y_{k}\right)=0
$$

regardless of the auxiliary vector, so we can calibrate with $\mathbf{t}_{\mathbf{O P}}(t)=a_{M}$ with dimension 1 .

Secondly, if we suppose that $D_{t}=\emptyset$, then $B_{t}=\emptyset$ and $A_{t}=A_{M}$. In this case, following ${ }^{3}$ the optimal auxiliary vector $\mathbf{t}_{\mathbf{O P}}(t)=\left(a_{1}, a_{2}, \ldots, a_{M}\right)$.

Since $A_{t}=Z_{t} \cap F_{t}=A_{M}$, if we suppose that $Z_{t}=A_{t}=A_{M}$ then $t>y_{k} \forall k \in U$ and this case is like the case where $t=y_{\text {max }}$ and although the optimal auxiliary vector is $\mathbf{t}_{\mathbf{O P}}(t)=\left(a_{1}, a_{2}, \ldots, a_{M}\right)$, we can reach the minimum value of $Q_{\gamma}(t)$ with the auxiliary vector $\gamma=\left(a_{M}\right)$.

On the other hand, if we consider that $Z_{t}=\emptyset$ and $F_{t}=A_{M}$ there is not reduction in the optimal auxiliary vector $\mathbf{t}_{\mathbf{O P}}(t)$ (see Appendix A.2.

Next, if we suppose that $Z_{t} \neq A_{t}=A_{M}$ and $F_{t} \neq A_{t}=A_{M}$ then there is a set $I_{F_{t}}=\left\{j_{1}, j_{2}, \ldots, j_{l}\right\} \subseteq\{1,2, \ldots M\}$ such that $a_{j_{i}} \in F_{t}$ and therefore $q_{j_{i}}^{t} \neq 0$ for $i=1,2, \ldots, l$.

Now, if we consider that $j_{1}>1 ; j_{i}-1>j_{(i-1)}$ for all $i=2, \ldots l$ and $j_{l}<M$; then for $h=1, \ldots j_{1}-1 ; q_{h}^{t}=0$ and we have:

$$
\begin{gathered}
K_{t}\left(a_{1}\right)=\sum_{k \in U} \Delta\left(a_{1}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{1}\right) \\
\vdots \\
K_{t}\left(a_{\left(j_{1}-1\right)}\right)=\sum_{k \in U} \Delta\left(a_{j_{1}-1}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{j_{1}-1}\right)
\end{gathered}
$$

Similarly, for $j_{i}, \ldots j_{i+1}-1$ with $i=1,2, \ldots l-1$ we have:

$$
\begin{gathered}
K_{t}\left(a_{j_{i}}\right)=\sum_{k \in U} \Delta\left(a_{j_{i}}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{j_{i}}\right)-\sum_{h=1}^{i} q_{j_{h}}^{t} \\
\vdots \\
K_{t}\left(a_{\left(j_{(i+1)}-1\right)}\right)=\sum_{k \in U} \Delta\left(a_{\left(j_{(i+1)}-1\right)}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{\left(j_{(i+1)}-1\right)}\right)-\sum_{h=1}^{i} q_{j_{h}}^{t}
\end{gathered}
$$

and finally, for $j_{l}, \ldots, M$

$$
\begin{gathered}
K_{t}\left(a_{j_{l}}\right)=\sum_{k \in U} \Delta\left(a_{j_{l}}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{j_{l}}\right)-\sum_{h=1}^{l} q_{j_{h}}^{t} \\
\vdots \\
K_{t}\left(a_{M}\right)=\sum_{k \in U} \Delta\left(a_{M}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{M}\right)-\sum_{h=1}^{l} q_{j_{h}}^{t}=N F_{y}(t) .
\end{gathered}
$$

The minimum of $Q_{t}(\gamma)$ reached at the optimum auxiliary vector $\mathbf{t}_{\mathbf{o P}}(t)$ is given by:

$$
\begin{gather*}
Q_{t}\left(\mathbf{t}_{\mathbf{O P}}(t)\right)=\left(N F_{y}(t)\right)^{2}-\sum_{j=1}^{M} \frac{\left(K_{t}\left(a_{j}-K_{t}\left(a_{j-1}\right)\right)\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}= \\
=\left(N F_{y}(t)\right)^{2}-N^{2} \cdot \sum_{\substack{j=1 \\
j \notin\left\{j_{1}, \ldots, j_{l}\right\}}}^{M} \frac{\left.\left(F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)\right)\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}-\sum_{j \in\left\{j_{1}, \ldots, j_{l}\right\}} \frac{\left.\left(N F_{g}\left(a_{j}\right)-N F_{g}\left(a_{j-1}\right)\right)-q_{j}^{t}\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)} \\
\left.=\left(N F_{y}(t)\right)^{2}-N^{2} \cdot \sum_{j=1}^{M}\left(F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)\right)\right)-N^{2} \cdot \sum_{j \in I_{F_{t}}}\left(F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)\right)+2 N \sum_{j \in I_{F_{t}}} q_{j}^{t}-\sum_{j \in I_{F_{t}}} \frac{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}{F_{j}}= \\
\left.=\left(N F_{y}(t)\right)^{2}-N^{2} \cdot \sum_{j=1}^{M}\left(F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)\right)\right)+2 N \sum_{j \in I_{F_{t}}} q_{j}^{t}-\sum_{j \in I_{F_{t}}}^{\sum_{g}} \frac{\left(q_{j}^{t}\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}= \\
=\left(N F_{y}(t)\right)^{2}-N^{2}+2 N \sum_{j \in I_{F_{t}}} q_{j}^{t}-\sum_{j \in I_{F_{t}}} \frac{\left(q_{j}^{t}\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)} . \tag{19}
\end{gather*}
$$

The same value can be reached with the auxiliary vector $\gamma=\left(a_{\left(j_{1}-1\right)}, a_{j_{1}}, \ldots a_{\left(j_{l}-1\right)}, a_{j_{l}}, a_{M}\right)$. To see it, if we set $a_{j 0}$ such as $F_{g}\left(a_{j 0}\right)=0$ and we replace the vector $\gamma$ in 15, we have:

$$
\begin{gathered}
Q_{t}(\gamma)=\left(N \cdot F_{y}(t)\right)^{2}-\sum_{h=1}^{l} \frac{\left(N F_{g}\left(a_{\left(j_{h}-1\right)}\right)-N F_{g}\left(a_{j_{(h-1)}}\right)\right)^{2}}{F_{g}\left(a_{\left(j_{h}-1\right)}\right)-F_{g}\left(a_{j_{(h-1)}}\right)}-\sum_{h=1}^{l} \frac{\left.\left(N F_{g}\left(a_{j_{h}}\right)-N F_{g}\left(a_{\left(j_{h}-1\right)}\right)-q_{j_{h}}^{t}\right)\right)^{2}}{F_{g}\left(a_{j_{h}}\right)-F_{g}\left(a_{\left(j_{h}-1\right)}\right)} \\
-\frac{\left(N F_{g}\left(a_{(M}\right)-N F_{g}\left(a_{j_{l}}\right)\right)^{2}}{F_{g}\left(a_{M}\right)-F_{g}\left(a_{j_{l}}\right)}=\left(N \cdot F_{y}(t)\right)^{2}-N^{2} \sum_{h=1}^{l} F_{g}\left(a_{\left(j_{h}-1\right)}\right)-F_{g}\left(a_{j_{(h-1)}}\right)-N^{2} \sum_{h=1}^{l} F_{g}\left(a_{j_{h}}\right)-F_{g}\left(a_{\left(j_{h}-1\right)}\right) \\
+2 N \sum_{h=1}^{l} q_{j_{h}}^{t}-\sum_{h=1}^{l} \frac{\left(q_{j_{h}}^{t}\right)^{2}}{F_{g}\left(a_{j_{h}}\right)-F_{g}\left(a_{\left(j_{h}-1\right)}\right)}-N^{2}\left(F_{g}\left(a_{M}\right)-F_{g}\left(a_{j_{l}}\right)\right) \\
=2 N \sum_{h=1}^{l} q_{j_{h}}^{t}-\sum_{h=1}^{l} \frac{F_{g}\left(a_{j_{h}}\right)-F_{g}\left(a_{\left(j_{h}-1\right)}\right)}{2}
\end{gathered}
$$

and the auxiliary vector $\gamma$, with less dimension than $\mathbf{t}_{\mathbf{O P}}(t)$, attain the minimum of $Q_{t}(\gamma)$.
Previously, we suppose that $a_{j_{1}}>a_{1}$ and $a_{j_{l}}<a_{M}$. If $a_{j_{1}}=a_{1}$ then it is easy to see that the minimum can be obtain at $\gamma=\left(a_{1}, a_{\left(j_{2}-1\right)}, a_{j_{2}}, \ldots a_{\left(j_{l}-1\right)}, a_{j_{l}}, a_{M}\right)$ that has less dimension than in the case $a_{j_{1}}>a_{1}$. In a similar way, if $a_{j_{l}}=a_{M}$ the minimum can be attained at $\gamma=\left(a_{j_{1}}, a_{\left(j_{1}-1\right)}, a_{j_{2}}, \ldots a_{\left(j_{l}-1\right)}, a_{M}\right)$.
Finally, we have assumed that $j_{i}-1>j_{(i-1)}$ for all $i=2, \ldots l$. If there is a $h \in\{1,2, \ldots l\}$ that $j_{h}-1=j_{(h-1)}$ it is easy to see that the minimum value of $Q_{t}(\gamma)$ is reached at $\gamma=\left(a_{\left(j_{1}-1\right)}, a_{j_{1}}, \ldots a_{j_{(h-1)}}, a_{j_{h}}, \ldots, a_{\left(j_{l}-1\right)}, a_{j_{l}}, a_{M}\right)$ with less dimension than in the case $j_{i}-1>j_{(i-1)}$ for all $i=2, \ldots l$. Therefore, if $D_{t}=\emptyset$ we can reduce the optimal dimension when $F_{t} \neq A_{t}=A_{M}$.

Next, we consider the case where $D_{t} \neq \emptyset$ and $D_{t} \neq A_{M}$. Because $A_{t}=A_{M}-D_{t}$, we have $A_{t} \neq \emptyset$ and $A_{t} \neq A_{M}$. Therefore:

$$
A_{t}=\left\{a_{1}^{t}, a_{2}^{t}, \ldots, a_{M_{t}}^{t}\right\}=\left\{a_{f_{1}^{t}}, a_{f_{2}^{t}}, \ldots, a_{f_{M_{t}}^{t}}\right\}
$$

where $\left\{f_{1}^{t}, f_{2}^{t}, \ldots, f_{M_{t}}^{t}\right\} \subseteq\{1,2, \ldots M\}$.
In this case, if we suppose that $B_{t}=\emptyset$ then $f_{1}^{t}=1, \ldots, f_{M_{t}}^{t}=M_{t}$ and

$$
A_{t}=\left\{a_{1}, a_{2}, \ldots a_{M_{t}}\right\} \quad ; \quad D_{t}=\left\{a_{M_{t}+1}, \ldots, a_{M}\right\}
$$

To see it, if we suppose that $f_{1}^{t}>1$ then $a_{f_{1}^{t}}>a_{\left(f_{1}^{t}-1\right)} \geq a_{1}$ and consequently the set

$$
U_{1}=\left\{l \in U: g_{l}<a_{1}^{t}\right\}=\left\{l \in U: g_{l}<a_{f_{1}^{t}}\right\} \neq \emptyset
$$

and $b_{1}^{t}=a_{\left(f_{1}^{t}-1\right)}$. Thus, $B_{t} \neq \emptyset$ (Contradiction). As a consequence, $a_{f_{1}^{t}}=a_{1}$.

If we suppose that for $i \in\left\{2, \ldots, M_{t}\right\}$ such as $f_{(i-1)}^{t}=i-1$ and we suppose that $f_{i}^{t}>i$ then $a_{f_{(i-1)}^{t}}=a_{(i-1)}$ and $a_{f_{i}^{t}}>a_{i}>$ $a_{(i-1)}=a_{f_{(i-1)}^{t}}$. The set $U_{i}$ is given by:

$$
U_{i}=\left\{l \in U: a_{(i-1)}^{t}<g_{l}<a_{i}^{t}\right\}=\left\{l \in U: a_{f_{(i-1)}}<g_{l}<a_{f_{i}}\right\} \neq \emptyset
$$

and $b_{i}^{t}=a_{\left(f_{i}^{t}-1\right)}$. Thus, $B_{t} \neq \emptyset$ (Contradiction again). As a consequence, if $f_{(i-1)}^{t}=i-1$ implies that $f_{i}^{t}=i$ for $i \in\left\{2, \ldots, M_{t}\right\}$ and we have:

$$
A_{t}=\left\{a_{1}, a_{2}, \ldots a_{M_{t}}\right\}
$$

If $M_{t}=M$ it is clear that $A_{t}=A_{M}$ and $D_{t}=\emptyset$ (Contradiction again). Therefore, $M_{t}<M$ and

$$
D_{t}=A_{M}-A_{t}=\left\{a_{\left(M_{t}+1\right)}, a_{\left(M_{t}+2\right)}, \ldots a_{M}\right\}
$$

The optimal auxiliary vector is given by $\mathbf{t}_{\mathbf{O P}}(t)=\left(a_{1}, \ldots, a_{M_{t}}\right)$ and in a similar way that in the previous cases, we can proof that if $F_{t}=A_{t}$, there is not a reduction in the optimal dimension. If $Z_{t}=A_{t}$, then we can attain the minimum of $Q_{t}(\gamma)$ at $\gamma=\left(a_{M_{t}}\right)$. Finally, if $Z_{t} \neq A_{t}$ and $F_{t} \neq A_{t}$, and we suppose that

$$
F_{t}=\left\{a_{j_{1}}, a_{j_{2}}, \ldots a_{j_{l}}\right\}
$$

we can reached the minimum value of $Q_{\gamma}(t)$ at $\gamma=\left(a_{\left(j_{1}-1\right)}, a_{j_{1}}, \ldots, a_{\left(j_{l}-1\right)}, a_{j_{l}}, a_{M_{t}}\right)$.

Next, under the assumption $D_{t} \neq \emptyset$ and $D_{t} \neq A_{M}$, we consider $B_{t} \neq \emptyset$. If we assume that $b_{h}^{t}$ exists for $h=1, \ldots, M_{t}$, following the proposal ${ }^{3}$, the optimal auxiliary vector $\mathbf{t}_{O P T}(t)$ is given by 12 and there is not a possible reduction in the dimension. On the other hand, if there are some values $j_{1}^{t}, j_{2}^{t}, \ldots j_{p_{t}}^{t} \in\left\{1, \ldots, M_{t}\right\}$; such as $b_{j_{h}}^{t}$ does not exits for $h=1,2, \ldots p_{t}$ with $p_{t} \leq M_{t}$ and $j_{h}^{t} \neq j_{q}^{t}$ if $h \neq q$, the optimal dimension is given by $P=2 M_{t}-p_{t}$ and the optimal auxiliary vector $\mathbf{t}_{\mathbf{O P}}$ is given by 13 . Analogously, there are $p_{1}, p_{2}, \ldots, p_{l_{t}} \in\left\{1,2, \ldots, M_{t}\right\}$ such as $b_{f_{p_{h}}}^{t}$ exists and following $^{3}$ there is not a reduction between the points $b_{f_{p_{h}}}^{t}$ and $a_{f_{p_{h}}}^{t}$. Alternatively, the optimal auxiliary vector $\mathbf{t}_{\mathbf{O P}}(t)$ can be expressed as follows

$$
\begin{equation*}
\mathbf{t}_{\mathbf{O P}}(t)=\left(a_{f_{1}^{t}}, \ldots, a_{\left.f_{\left(p_{1}-1\right)}^{t}\right)}, b_{f_{p_{1}}^{t}}, a_{f_{p_{1}}^{t}}, a_{f_{\left(p_{1}+1\right)}^{t}}, \ldots, a_{f_{\left(p_{h}-1\right)}^{t}}, b_{f_{p_{h}}^{t}}, a_{f_{p_{h}}^{t}}, a_{f_{\left(p_{h}+1\right)}^{t}}, \ldots, a_{f_{M_{t}}^{t}}\right) \tag{20}
\end{equation*}
$$

If we suppose that $p_{1}=1$ then the value $b_{f_{1}}^{t}$ exists and consequently:

$$
U_{1}=\left\{l \in U: g_{l}<a_{1}^{t}\right\}=\left\{l \in U: g_{l}<a_{f_{1}^{t}}\right\} \neq \emptyset
$$

then, $a_{f_{1}^{t}}>a_{1}$ and $f_{1}^{t}>1$. As a consequence $a_{i} \notin A_{t}$ with $i=1, \ldots f_{1}^{t}-1$ and $\left\{a_{1}, \ldots a_{\left(f_{1}^{t}-1\right)}\right\} \subseteq D_{t}$ which implies that $b_{f_{1}^{t}}=a_{\left(f_{1}^{t}-1\right)}$. In this case, we have:

$$
\begin{gathered}
K_{t}\left(b_{f_{1}^{t}}\right)=\sum_{k \in U} \Delta\left(b_{f_{1}^{t}}-g_{k}\right) \Delta\left(t-y_{k}\right)=0 \\
K_{t}\left(a_{f_{1}^{t}}\right)=\sum_{k \in U} \Delta\left(a_{f_{1}^{t}}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{f_{1}^{t}}\right)-\sum_{h=1}^{f_{1}^{t}-1} r_{h}-q_{1}^{t}=N F_{g}\left(a_{f_{1}^{t}}\right)-H_{1}^{t}
\end{gathered}
$$

and following ${ }^{3}$ there is no possibility to reduce the number of points here.
On the other hand, if we suppose that $p_{1}>1$, for $i \in\left\{1,2, \ldots p_{1}-1\right\}$ the value $b_{f_{i}}^{t}$ does not exist and:

$$
\begin{gathered}
U_{1}=\left\{l \in U: g_{l}<a_{1}^{t}\right\}=\left\{l \in U: g_{l}<a_{f_{1}^{t}}\right\}=\emptyset \\
U_{i}=\left\{l \in U: a_{(i-1)}^{t}<g_{l}<a_{i}^{t}\right\}=\left\{l \in U: a_{f_{(i-1)}^{t}}<g_{l}<a_{f_{i}^{t}}\right\}=\emptyset \text { for } i=2, \ldots, p_{1}-1
\end{gathered}
$$

therefore $a_{f_{1}^{t}}=a_{1}$ and $a_{f_{i}^{t}}=a_{\left(f_{(i-1)}+1\right)}$ for $i=2, \ldots, p_{1}-1$ and then:

$$
a_{f_{1}^{t}}=a_{1} ; a_{f_{2}^{t}}=a_{2} ; \ldots ; a_{f_{\left(p_{1}-1\right)}^{t}}=a_{\left(p_{1}-1\right)} .
$$

Thus, if $p_{1}>1$ we have $\left\{a_{1}, \ldots, a_{\left(p_{1}-1\right)}\right\} \subseteq A_{t}$ and consequently the optimal auxiliary vector $\mathbf{t}_{\mathbf{O P}}(t)$ given by 20 can be expressed as follows:

$$
\mathbf{t}_{\mathbf{O P}}(t)=\left(a_{1}, \ldots, a_{\left(p_{1}-1\right)}, b_{f_{p_{1}}^{t}}, a_{f_{p_{1}}^{t}}, \mathbf{t}_{R}\right)=
$$

where $\mathbf{t}_{R}$ denotes

$$
\mathbf{t}_{R}=\left(a_{f_{\left(p_{1}+1\right)}^{t}}, \ldots, a_{\left.f_{\left(p_{h}-1\right)}^{t}\right)}, b_{f_{p_{h}}^{t}}, a_{f_{p_{h}}^{t}}, a_{f_{\left(p_{h}+1\right)}^{t}}, \ldots, a_{f_{M_{t}}^{t}}\right)
$$

On the contrary, for $p_{1}$ we have

$$
\begin{gathered}
U_{p_{1}}=\left\{l \in U: a_{\left(p_{1}-1\right)}^{t}<g_{l}<a_{p_{1}}^{t}\right\}=\left\{l \in U: a_{f_{\left(p_{1}-1\right)}}<g_{l}<a_{f_{p_{1}}^{t}}\right\}= \\
=\left\{l \in U: a_{\left(p_{1}-1\right)}<g_{l}<a_{f_{p_{1}}^{t}}\right\} \neq \emptyset
\end{gathered}
$$

and then $a_{f_{p_{1}}}>a_{p_{1}}$ and $f_{p_{1}}^{t}>p_{1}$ which implies that

$$
\left\{a_{p_{1}}, a_{\left(p_{1}+1\right)}, \ldots a_{\left(f_{p_{1}}^{t}-1\right)}\right\} \subseteq D_{t}
$$

We consider the following sets:

$$
\begin{gather*}
A_{p_{1}}=\left\{a_{1}, \ldots, a_{\left(p_{1}-1\right)}\right\}  \tag{21}\\
Z_{p_{1}}=\left\{a_{i} \in A_{p_{1}}: q_{i}^{t}=0\right\}  \tag{22}\\
F_{p_{1}}=\left\{a_{i} \in A_{p_{1}}: 0<q_{i}^{t}<r_{i}\right\} \tag{23}
\end{gather*}
$$

Similarly to the previous cases, if we suppose that $Z_{p_{1}}=A_{p_{1}}$ and $F_{p_{1}}=\emptyset$, it is easy to see that we can delete $a_{i}$ for $i=1,2, \ldots p_{1}-2$ from the optimal auxiliary vector $\mathbf{t}_{\mathbf{O P}}(t)$ and we can calibrate only with the value $a_{\left(p_{1}-1\right)}$. To see it, it is clear that:

$$
\begin{gathered}
K_{t}\left(a_{1}\right)=\sum_{k \in U} \Delta\left(a_{1}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{1}\right) \\
\vdots \\
K_{t}\left(a_{\left(p_{1}-1\right)}\right)=\sum_{k \in U} \Delta\left(a_{j_{1}-1}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{\left(p_{1}-1\right)}\right) .
\end{gathered}
$$

Because $\left\{a_{p_{1}}, a_{\left(p_{1}+1\right)}, \ldots a_{\left(f_{p_{1}}-1\right)}\right\} \subseteq D_{t}$ it is easy to see that:

$$
\begin{gathered}
K_{t}\left(b_{f_{p_{1}}^{t}}\right)=\sum_{k \in U} \Delta\left(b_{f_{p_{1}}^{t}}-g_{k}\right) \Delta\left(t-y_{k}\right)=K_{t}\left(a_{\left(p_{1}-1\right)}\right)=N F_{g}\left(a_{\left(p_{1}-1\right)}\right) \\
K_{t}\left(a_{f_{p_{1}}}\right)=\sum_{k \in U} \Delta\left(a_{f_{p_{1}}^{t}}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{f_{p_{1}}^{t}}\right)-\sum_{h=p_{1}}^{f_{p_{1}}-1} r_{h}-q_{f_{p_{1}}}^{t}=N F_{g}\left(a_{f_{p_{1}}}\right)-H_{p_{1}}^{t} .
\end{gathered}
$$

The minimum value of $Q_{t}(\gamma)$ at $\mathbf{t}_{\mathbf{O P}}(t)$ can be expressed as follows:

$$
\begin{gathered}
Q_{t}\left(\mathbf{t}_{\mathbf{O P}}(t)\right)=Q_{t}\left(a_{1}, \ldots, a_{\left(p_{1}-1\right)}, b_{f_{p_{1}}}, a_{f_{p_{1}}}, \mathbf{t}_{R}\right)= \\
=Q_{t}\left(\mathbf{t}_{R}\right)-\sum_{j=1}^{p_{1}-1} \frac{\left(K_{t}\left(a_{j}\right)-K_{t}\left(a_{j-1}\right)\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}-\frac{\left(K_{t}\left(b_{f_{p_{1}}^{t}}\right)-K_{t}\left(a_{\left(p_{1}-1\right)}\right)\right)^{2}}{F_{g}\left(b_{f_{p_{1}}^{t}}\right)-F_{g}\left(a_{\left(p_{1}-1\right)}\right)}-\frac{\left(K_{t}\left(a_{f_{p_{1}}}\right)-K_{t}\left(b_{f_{p_{1}}^{t}}\right)\right)^{2}}{F_{g}\left(a_{f_{p_{1}}^{t}}\right)-F_{g}\left(b_{f_{p_{1}}^{t}}\right)}= \\
=Q_{t}\left(\mathbf{t}_{R}\right)-N^{2} \sum_{j=1}^{p_{1}-1}\left(F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)\right)-\frac{\left(N F_{g}\left(a_{f_{p_{1}}^{t}}\right)-N F_{g}\left(a_{\left(p_{1}-1\right)}\right)-H_{p_{1}}^{t}\right)^{2}}{F_{g}\left(a_{f_{p_{1}}}\right)-F_{g}\left(b_{f_{p_{1}}^{\prime}}\right)}= \\
=Q_{t}\left(\mathbf{t}_{R}\right)-N^{2} F_{g}\left(a_{\left(p_{1}-1\right)}\right)-\frac{\left(N F_{g}\left(a_{f_{p_{1}}}\right)-N F_{g}\left(a_{\left(p_{1}-1\right)}\right)-H_{p_{1}}^{t}\right)^{2}}{F_{g}\left(a_{f_{p_{1}}}\right)-F_{g}\left(b_{f_{p_{1}}^{t}}\right)} .
\end{gathered}
$$

If we calibrate with the auxiliary vector $\gamma=\left(a_{\left(p_{1}-1\right)}, b_{f_{p_{1}}}, a_{f_{p_{1}}}, \mathbf{t}_{R}\right)$, we obtain the same value:

$$
\begin{gathered}
Q_{t}(\gamma)=Q_{t}\left(a_{\left(p_{1}-1\right)}, b_{f_{p_{1}}}, a_{f_{p_{1}}}, \mathbf{t}_{R}\right)= \\
=Q_{t}\left(\mathbf{t}_{R}\right)-\frac{\left.\left(K_{t}\left(a_{\left(p_{1}-1\right)}\right)\right)\right)^{2}}{F_{g}\left(a_{\left(p_{1}-1\right)}\right)}-\frac{\left(K_{t}\left(b_{f_{p_{1}}^{t}}\right)-K_{t}\left(a_{\left(p_{1}-1\right)}\right)\right)^{2}}{F_{g}\left(b_{f_{p_{1}}^{t}}\right)-F_{g}\left(a_{\left(p_{1}-1\right)}\right)}-\frac{\left(K_{t}\left(a_{f_{p_{1}}^{t}}\right)-K_{t}\left(b_{f_{p_{1} t}}\right)\right)^{2}}{F_{g}\left(a_{f_{p_{1}}^{t}}\right)-F_{g}\left(b_{f_{p_{1}}^{t}}\right)}= \\
=Q_{t}\left(\mathbf{t}_{R}\right)-N^{2} F_{g}\left(a_{\left(p_{1}-1\right)}\right)-\frac{\left(N F_{g}\left(a_{f_{p_{1}}^{t}}\right)-N F_{g}\left(a_{\left(p_{1}-1\right)}\right)-H_{p_{1}}^{t}\right)^{2}}{F_{g}\left(a_{f_{p_{1}}}\right)-F_{g}\left(b_{f_{p_{1}}}\right)} .
\end{gathered}
$$

Thus, if $Z_{p_{1}}=A_{p_{1}}$ we can reduce the dimension of the auxiliary vector to attain the minimum value of $Q_{t}(\gamma)$.
Now, if we suppose that $F_{p_{1}}=A_{p_{1}}$ and $Z_{p_{1}}=\emptyset$, as in the previous cases, we cannot reduce the dimension of the subvector $\left(a_{1}, \ldots, a_{\left(p_{1}-1\right)}\right)$ in the auxiliary vector $\left(\mathbf{t}_{\mathbf{O P}}(t)\right)$.

Next, if we suppose that $Z_{p_{1}} \neq A_{p_{1}}$ and $F_{p_{1}} \neq A_{p_{1}}$ then there is a set $I_{F_{p_{1}}}=\left\{j_{1}^{1}, j_{2}^{2}, \ldots, j_{l_{1}}^{1}\right\} \subseteq\left\{1,2, \ldots p_{1}-1\right\}$ such that $a_{j_{h}^{1}} \in F_{p_{1}}$ and therefore $q_{j_{h}^{1}}^{t} \neq 0$ for $h=1,2, \ldots, l_{1}$.

Now, if we consider that $j_{1}^{1}>1 ; j_{h}^{1}-1>j_{(h-1)}^{1}$ for all $h=2, \ldots l_{1}$ and $j_{l_{1}}^{1}<p_{1}-1$; then for $v=1, \ldots j_{1}^{1}-1 ; q_{v}^{t}=0$ and we have:

$$
\begin{gathered}
K_{t}\left(a_{1}\right)=\sum_{k \in U} \Delta\left(a_{1}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{1}\right) \\
\vdots \\
K_{t}\left(a_{\left(j_{1}^{1}-1\right)}\right)=\sum_{k \in U} \Delta\left(a_{j_{1}^{1}-1}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{j_{1}^{1}-1}\right)
\end{gathered}
$$

Similarly, for $j_{h}^{1}, \ldots j_{h+1}^{1}-1$ with $h=1,2, \ldots l_{1}-1$; we have:

$$
\left.\begin{array}{c}
K_{t}\left(a_{j_{h}}\right)=\sum_{k \in U} \Delta\left(a_{j_{h}^{1}}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{j_{h}^{1}}\right)-\sum_{v=1}^{h} q_{j_{v}^{1}}^{t} \\
\vdots \\
K_{t}\left(a_{\left(j_{(h+1)}^{1}\right.}-1\right)
\end{array}\right)=\sum_{k \in U} \Delta\left(a_{\left(j_{(h+1)}^{1}-1\right)}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{\left(j_{(h+1)}^{1}-1\right)}\right)-\sum_{v=1}^{h} q_{j_{v}^{1}}^{t}
$$

and finally, for $j_{l_{1}}^{1}, \ldots, p_{1}-1$

$$
\begin{gathered}
K_{t}\left(a_{j_{l_{1}}}\right)=\sum_{k \in U} \Delta\left(a_{j_{l_{1}}}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{j_{l_{1}}}\right)-\sum_{v=1}^{l_{1}^{1}} q_{j_{v}^{1}}^{t}=N F_{g}\left(a_{j_{l_{1}}}\right)-L_{1}^{t} \\
\vdots \\
K_{t}\left(a_{\left(p_{1}-1\right)}\right)=\sum_{k \in U} \Delta\left(a_{\left(p_{1}-1\right)}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{\left(p_{1}-1\right)}\right)-\sum_{v=1}^{l_{1}^{1}} q_{j_{v}^{1}}^{t}=N F_{g}\left(a_{\left(p_{1}-1\right)}\right)-L_{1}^{t} .
\end{gathered}
$$

Again, because $\left\{a_{p_{1}}, a_{\left(p_{1}+1\right)}, \ldots a_{\left(f_{p_{1}}-1\right)}\right\} \subseteq D_{t}$ we have:

$$
\begin{gathered}
K_{t}\left(b_{f_{p_{1}}^{t}}\right)=\sum_{k \in U} \Delta\left(b_{f_{p_{1}}^{t}}-g_{k}\right) \Delta\left(t-y_{k}\right)=K_{t}\left(a_{\left(p_{1}-1\right)}\right)=N F_{g}\left(a_{\left(p_{1}-1\right)}\right)-L_{1}^{t} \\
K_{t}\left(a_{f_{p_{1}}}\right)=\sum_{k \in U} \Delta\left(a_{f_{p_{1}}^{t}}-g_{k}\right) \Delta\left(t-y_{k}\right)=N F_{g}\left(a_{f_{p_{1}}}\right)-L_{1}^{t}-\sum_{h=p_{1}}^{f_{p_{1}}-1} r_{h}-q_{f_{p_{1}}^{t}}^{t}=N F_{g}\left(a_{f_{p_{1}}}\right)-L_{1}^{t}-H_{p_{1}}^{t}
\end{gathered}
$$

The minimum of $Q_{t}(\gamma)$ at $\left(\mathbf{t}_{\mathbf{O P}}(t)\right)$ is given by:

$$
\begin{gathered}
Q_{t}\left(\mathbf{t}_{\mathbf{O P}}(t)\right)=Q_{t}\left(a_{1}, \ldots, a_{\left(p_{1}-1\right)}, b_{f_{p_{1}}^{t}}, a_{f_{p_{1}}^{t}}, \mathbf{t}_{R}\right)= \\
=Q_{t}\left(\mathbf{t}_{R}\right)-\sum_{j=1}^{p_{1}-1} \frac{\left(K_{t}\left(a_{j}\right)-K_{t}\left(a_{j-1}\right)\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}-\frac{\left(K_{t}\left(b_{f_{p_{1}}^{t}}\right)-K_{t}\left(a_{\left(p_{1}-1\right)}\right)\right)^{2}}{F_{g}\left(b_{f_{p_{1}}^{t}}\right)-F_{g}\left(a_{\left(p_{1}-1\right)}\right)}-\frac{\left(K_{t}\left(a_{f_{p_{1}}^{t}}\right)-K_{t}\left(b_{f_{p_{1}}^{t}}\right)\right)^{2}}{F_{g}\left(a_{f_{p_{1}}^{t}}\right)-F_{g}\left(b_{f_{p_{1}}^{t}}\right)}= \\
=Q_{t}\left(\mathbf{t}_{R}\right)-N^{2} \sum_{j=1}^{p_{1}-1}\left(F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)\right)+2 N \cdot L_{1}^{t}-\sum_{v=1}^{l_{1}^{1}} \frac{\left(q_{j_{v}^{\prime}}^{t}\right)^{2}}{F_{g}\left(a_{j_{v}}\right)-F_{g}\left(a_{\left(j_{1}^{v}-1\right)}\right)}-\frac{\left(N F_{g}\left(a_{f_{p_{1}}^{t}}\right)-N F_{g}\left(a_{\left(p_{1}-1\right)}\right)-H_{p_{1}}^{t}\right)^{2}}{F_{g}\left(a_{f_{p_{1}}}\right)-F_{g}\left(b_{f_{p_{1} t}}\right)}= \\
=Q_{t}\left(\mathbf{t}_{R}\right)-N^{2} F_{g}\left(a_{\left(p_{1}-1\right)}\right)+2 N \cdot L_{1}^{t}-\sum_{v=1}^{l_{1}^{1}} \frac{\left(q_{j_{v}^{1}}^{t}\right)^{2}}{F_{g}\left(a_{j_{v}}\right)-F_{g}\left(a_{\left(j_{v}^{1}-1\right)}\right)}-\frac{\left(N F_{g}\left(a_{f_{p_{1}}^{t}}\right)-N F_{g}\left(a_{\left(p_{1}-1\right)}\right)-H_{p_{1}}^{t}\right)^{2}}{F_{g}\left(a_{f_{p_{1} t}}\right)-F_{g}\left(b_{f_{p_{1} t}}\right)} .
\end{gathered}
$$

If we calibrate with the following auxiliary vector:

$$
\gamma=\left(a_{\left(j_{1}^{1}-1\right)}, a_{j_{1}^{1}}, \ldots, a_{\left(j_{l_{1}}^{1}-1\right)}, a_{j_{l_{1}}^{1}}, a_{\left(p_{1}-1\right)}, b_{f_{p_{1}}^{t}}, a_{f_{p_{1}}^{\prime}}, \mathbf{t}_{R}\right)
$$

in a similar way to the previous cases, it is easy to see that:

$$
Q_{t}\left(a_{\left(j_{1}^{1}-1\right)}, a_{j_{1}^{1}}, \ldots, a_{\left(j_{l_{1}}^{1}-1\right)}, a_{j_{l_{1}}}, a_{\left(p_{1}-1\right)}, b_{f_{p_{1}}^{t}}, a_{f_{p_{1}}^{t}}, \mathbf{t}_{R}\right)=
$$

$$
=Q_{t}\left(\mathbf{t}_{R}\right)-N^{2} F_{g}\left(a_{\left(p_{1}-1\right)}\right)+2 N \cdot L_{1}^{t}-\sum_{v=1}^{l_{1}^{1}} \frac{\left(q_{j_{v}^{1}}^{t}\right)^{2}}{F_{g}\left(a_{j_{v}^{1}}\right)-F_{g}\left(a_{\left(j_{v}^{1}-1\right)}\right)}-\frac{\left(N F_{g}\left(a_{f_{p_{1}}}\right)-N F_{g}\left(a_{\left(p_{1}-1\right)}\right)-H_{p_{1}}^{t}\right)^{2}}{F_{g}\left(a_{f_{p_{1}}^{t}}\right)-F_{g}\left(b_{f_{p_{1}}^{t}}\right)}
$$

Then, if $Z_{p_{1}} \neq A_{p_{1}}$ and $F_{p_{1}} \neq A_{p_{1}}$ again we can reduce the dimension of the optimal vector $\mathbf{t}_{\mathbf{O P}}(t)$.
Finally, we have assumed that $j_{1}^{1}>1 ; j_{h}^{1}-1>j_{(h-1)}^{1}$ for all $h=2, \ldots l_{1}$ and $j_{l_{1}}^{1}<p_{1}-1$. If there is a $h \in\{1,2, \ldots l\}$ that $j_{h}^{1}-1=j_{(h-1)}^{1}$ it is easy to see that the minimum value of $Q_{t}(\gamma)$ is reached at

$$
\gamma=\left(a_{\left(j_{1}^{1}-1\right)}, a_{j_{1}^{1}}, \ldots a_{j_{(h-1)}^{1}}, a_{j_{h}^{1}}, \ldots, a_{\left(j_{l_{1}}^{1}-1\right)}, a_{j_{l_{1}}}, a_{\left(p_{1}-1\right)}, b_{f_{p_{1}}^{t}}, a_{f_{p_{1}}^{t}}, \mathbf{t}_{R}\right)
$$

with less dimension than in the case $j_{h}^{1}-1>j_{(h-1)}^{1}$ for all $h=2, \ldots l_{1}$. Similarly, if we suppose that $j_{1}^{1}=1$ then the minimum is obtained with

$$
\gamma=\left(a_{j_{1}^{1}}, a_{\left(j_{2}^{1}-1\right)}, a_{j_{2}^{1}}, \ldots a_{j_{(h-1)}^{1}}, a_{j_{h}^{1}}, \ldots, a_{\left(j_{l_{1}}^{1}-1\right)}, a_{j_{l_{1}}^{1}}, a_{\left(p_{1}-1\right)}, b_{f_{p_{1}}^{t}}, a_{f_{p_{1}}^{t}}, \mathbf{t}_{R}\right)
$$

again, with less dimension than in the case $j_{1}^{1}>1$. In a similar way, if $j_{l_{1}}^{1}=p_{1}-1$, the minimum of $Q_{t}(\gamma)$ is reached at

$$
\gamma=\left(a_{\left(j_{1}^{1}-1\right)}, a_{j_{1}^{1}}, \ldots a_{j_{(h-1)}^{1}}, a_{j_{h}^{1}}, \ldots, a_{\left(j_{l_{1}}^{1}-1\right)}, a_{j_{l_{1}}}, b_{f_{p_{1}}^{t}}, a_{f_{p_{1}}^{t}}, \mathbf{t}_{R}\right)
$$

again, with less dimension than in the case $j_{l_{1}}^{1}<p_{1}-1$.
Now, if we consider $p_{i}$ with $i=2, \ldots, l_{t}$, we can extend the analysis for the dimension reduction of the optimal auxiliary vector in a similar way to the case $p_{1}$. (see Appendix A.3).

## 4 | THE NEW OPTIMAL ESTIMATOR WITH THE NEW OPTIMAL VECTOR

In the previous section we have theoretically demonstrated that the optimal vector $\mathbf{t}_{\mathbf{O P T}}(t)$ proposed in ${ }^{33}$ can be reduced in its dimension and we can minimize the variance given by 8$]$ with a new optimal vector $\mathbf{t}_{\text {NEWOPT }}(t)$ of lower dimension. As with the original optimal vector $\mathbf{t}_{\mathbf{O P T}}(t)$, the new vector $\mathbf{t}_{\text {NEWOPT }}(t)$ depends on unknown population values and therefore needs to be estimated. For it, in the same way as in ${ }^{3}$, from the sample versions of the sets $A_{t}, B_{t}, C_{t}, D_{t}, Z_{t}$ and $F_{t}$ and the sample version of the function $Q_{t}(\gamma)$, an estimate $\widehat{\mathbf{t}}_{\text {NEWOPT }}(t)$ of the vector $\mathbf{t}_{\text {NEWOPT }}(t)$ can be obtained and we can define a new calibrated estimator $\widehat{F}_{C A L N E W O P T}(t)$ for the distribution function $F_{y}(t)$, given by:

$$
\begin{equation*}
\widehat{F}_{C A L N E W O P T}(t)=\widehat{F}_{Y H T}(t)+\left(F_{g}\left(\widehat{\mathbf{t}}_{\text {NEWOPT }}(t)\right)-\widehat{F}_{G H T}\left(\widehat{\mathbf{t}}_{\text {NEWOPT }}(t)\right)\right)^{\prime} \cdot \widehat{D}\left(\widehat{\mathbf{t}}_{\text {NEWOPT }}(t)\right) \tag{24}
\end{equation*}
$$

where

$$
\left.\widehat{D}\left(\hat{\mathbf{t}}_{\text {NEWOPT }}(t)\right)=T^{-1} \cdot \sum_{k \in s} d_{k} q_{k} \Delta \widehat{\mathbf{t}}_{\text {NEWOPT }}(t)-g_{k}\right) \Delta\left(t-y_{k}\right)
$$

## 5 | SIMULATION STUDY

In this section, a simulation study was conducted to compare the performance of the proposed optimal estimator with other alternative estimators for the distribution function $F(t)$. The simulation study was programmed in R software [version 4.1.0] and it was necessary to develop a new code to calculate the estimators included in the simulation study. The precision of the proposed new optimal calibration estimator $\widehat{F}_{C A L N E W O P T}(t)$ was compared with the following estimators, the Horvitz Thompson estimator, $\widehat{F}_{H T}$, the difference estimator ${ }^{31}, \widehat{F}_{D}(t)$, the ratio estimator ${ }^{31} \widehat{F}_{R}(t)$, the Chambers-Dunstan estimator ${ }^{32]} \widehat{F}_{C D}(t)$, the Rao-Kovar-Mantel estimator ${ }^{31]} \widehat{F}_{R K M}(t)$, the calibration estimator ${ }^{[2]}$ with $t_{1}=Q_{g}(0.5)$, the population median, as point for calibration, $\widehat{F}_{C A L}(t)$, the calibration estimator ${ }^{[2]}$ with three points $t_{1}=Q_{g}(0.25), t_{2}=Q_{g}(0.5)$ and $t_{3}=Q_{g}(0.75)$, the population quartiles, as points for calibration, $\widehat{F}_{C A L, 3}(t)$, the calibration estimator with one optimal point $\frac{18}{18}, \widehat{F}_{C A L M A X}(t)$ and finally the previous optimal calibration estimator ${ }^{[3]} \widehat{F}_{C A L O P T}(t)$.

Both real populations and simulated populations were considered for the simulation study. Specifically, we considered a real population included in The R Datasets Package called DNase that provides data collected from an ELISA assay for recombinant

DNase protein in rat serum with population size $N=176$. In addition, two simulated population called Simh and Simser were considered. The first one, Simh, is a population of size $N=5000$ generated from the following superpopulation model:

$$
y_{k}=8-7.82 / x_{k}+\epsilon_{k}
$$

where $x$ is a sample from a discrete uniform distribution in $\{1,2, \ldots, 100\}$ and $\epsilon_{k}$ 's are i.i.d. random variables from $N\left(0,0.5 / x_{k}\right)$.

The simulated population Simser is a population of size $N=5882$ generated from the following superpopulation model:

$$
y_{k}=x_{k}^{2}+\epsilon_{k}
$$

where $x$ is a sample from a discrete uniform distribution in $\{-100,-99, \ldots, 100\}$ and $\epsilon_{k}$ 's are i.i.d. random variables from $N(0,10)$.

For each population included in the simulation study, we drawn by simple random sampling without replacement 1000 samples of several sizes. For each sample, we estimated the distribution function $F(t)$ through all the estimators considered in the study at 11 different values of $t$, namely the quantiles $Q_{y}(\alpha)$ for $\alpha=0.1,0.2,0.25,0.3,0.4,0.5,0.6,0.7,0.75,0.8$ and 0.9 .

To measure the performance of each estimator included in the study, we considered the average relative bias (AVRB) and the average relative efficiency (AVRE), defined as follow:

$$
\operatorname{AVRB}(t)=\frac{1}{11} \sum_{q=1}^{11}\left|\operatorname{RB}\left(t_{q}\right)\right|, \quad \operatorname{AVRE}(t)=\frac{1}{11} \sum_{q=1}^{11} \operatorname{RE}\left(t_{q}\right)
$$

where RB and RE are defined as

$$
\begin{equation*}
\operatorname{RB}(t)=\frac{1}{B} \sum_{b=1}^{B} \frac{\widehat{F}(t)_{b}-F_{y}(t)}{F_{y}(t)} \quad \text { and } \quad \operatorname{RE}(t)=\frac{M S E[\widehat{F}(t)]}{M S E\left[\widehat{F}_{H T}(t)\right]} \tag{25}
\end{equation*}
$$

where $b$ indexes the $b$ th simulation run, $\widehat{F}(t)$ is an estimator for the distribution function, $M S E[\widehat{F}(t)]=B^{-1} \sum_{b=1}^{B}\left[\widehat{F}(t)_{b}-F_{y}(t)\right]^{2}$ is the empirical mean square error for $\widehat{F}(t)$ and $\operatorname{MSE}\left[\widehat{F}_{H T}(t)\right]$ is similarly defined for the Horvitz-Thompson estimator.

Given that the new estimator proposal $\widehat{F}_{C A L N E W O P T}(t)$ and the estimator $\widehat{F}_{C A L O P T}(t)$ are based on the minimization of 15 ,, it is possible that their behavior in terms of efficiency is similar and therefore it is necessary to analyze their behavior in greater detail. A reduced dimensionality in the auxiliary information set may reduce numerical issues in optimization procedures and also avoid the presence of unstable calibration weights (both negative weights and huge weights). Therefore, for each of the eleven estimation points $t_{q}$, we compared the dimension of the optimum auxiliary vector used in each estimators $\widehat{F}_{C A L O P T}(t)$ and $\widehat{F}_{C A L N E W O P T}(t)$. For it, we considered the mean dimension and the variance of the dimension:

$$
\operatorname{MD}\left(\widehat{F}\left(t_{q}\right)\right)=\frac{1}{B} \sum_{b=1}^{B} \operatorname{DIM}\left(\mathbf{t}_{\mathbf{q o p t}}\right), \quad \mathrm{VD}\left(\widehat{F}\left(t_{q}\right)\right)=\frac{1}{B} \sum_{b=1}^{B}\left(\operatorname{DIM}\left(\mathbf{t}_{\mathbf{q o p t}}\right)-\operatorname{MD} t_{q}\right)^{2}
$$

where $\widehat{F}$ can be $\widehat{F}_{C A L O P T}(t)$ or $\widehat{F}_{C A L N E W O P T}(t), \mathbf{t}_{\text {qopt }}$ denote the optimum auxiliary vector used with the point $t_{q}$ and DIM $\left(\mathbf{t}_{\text {qopt }}\right)$ denote the dimension of $\left(\mathbf{t}_{\mathbf{q o p t}}\right)$.

Additionally, because a limited number of variables may reduce the execution time to resolve the calibration procedure, we compared for each estimation point the execution time in calculating the estimators using the following measure:

$$
\operatorname{RT}(t)=\frac{\operatorname{TIME}\left(\widehat{F}_{C A L N E W O P T}(t)\right)}{\operatorname{TIME}\left(\widehat{F}_{C A L O P T}(t)\right)}
$$

where $\operatorname{TIME}\left(\widehat{F}_{C A L O P T}(t)\right)$ and $\operatorname{TIME}\left(\widehat{F}_{C A L N E W O P T}(t)\right)$ denote the running time for calculating $\widehat{F}_{C A L O P T}(t)$ and $\widehat{F}_{C A L N E W O P T}(t)$ respectively.

For the population DNase, Table 1 gives the values of AVRB and AVRE whereas Table 2 gives the values of MD; VD and RT. With respect to the results obtained for the bias and efficiency analysis, the estimators $\widehat{F}_{C A L O P T}(t)$ and $\widehat{F}_{C A L N E W O P T}(t)$ show the same behavior. Thus, both estimators present an adequate bias value, although for all sample sizes there are estimators that present a lower bias. In relation to efficiency, both estimators are clearly the most efficient. From results in Table 2 as expected, the estimator $\widehat{F}_{C A L N E W O P T}(t)$ always presents a smaller dimension of the auxiliary vector used, but this reduction is quite modest for all sample sizes and therefore the reduction obtained in the execution time is also quite modest. This may be because the set $F_{t}$ has a cardinal similar to the set $A_{t}$ or the set $B_{t}$ also has a cardinal similar to the set $A_{t}$. Consequently, the new
proposal $\widehat{F}_{C A L N E W O P T}(t)$ only achieves small reductions in the dimension of the auxiliary information used with respect to $\widehat{F}_{C A L O P T}(t)$, that produces slight improvements in execution time and it is not considerable enough to improve the asymptotic behavior ${ }^{[23]}$, although it does not deteriorate it either.

TABLE 1 Average relative bias (AVRB) and average relative efficiency (AVRE) of the estimators compared. Population: DNase.

|  |  |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | AVRB | AVRE | AVRB | AVRE | AVRB | AVRE | AVRB | AVRE |
|  | $n=30$ | $n=32$ |  | $n=35$ |  | $n=37$ |  |  |
| $\widehat{F}_{H T}$ | 0.0064 | 1 | 0.0063 | 1 | 0.0037 | 1 | 0.0030 | 1 |
| $\widehat{F}_{D}$ | 0.0136 | 0.4898 | 0.0219 | 0.4867 | 0.0173 | 0.4642 | 0.0153 | 0.4703 |
| $\widehat{F}_{R}$ | 0.0083 | 0.4412 | 0.0040 | 0.4343 | 0.0061 | 0.4247 | 0.0024 | 0.4335 |
| $\widehat{F}_{C D}$ | 0.1869 | 0.8930 | 0.1878 | 0.9150 | 0.1814 | 0.9441 | 0.1783 | 0.9749 |
| $\widehat{F}_{R K M}$ | 0.0032 | 0.4128 | 0.0103 | 0.4124 | 0.0063 | 0.4023 | 0.0055 | 0.4068 |
| $\widehat{F}_{C A L}$ | 0.0066 | 0.8575 | 0.0090 | 0.8377 | 0.0029 | 0.8881 | 0.0012 | 0.8436 |
| $\widehat{F}_{C A L 3}$ | 0.0046 | 0.3401 | 0.0035 | 0.3468 | 0.0053 | 0.3322 | 0.0015 | 0.3261 |
| $\widehat{F}_{C A L M A X}$ | 0.0057 | 0.2102 | 0.0079 | 0.2247 | 0.0052 | 0.1981 | 0.0050 | 0.1991 |
| $\widehat{F}_{C A L O P T}$ | 0.0047 | 0.1895 | 0.0059 | 0.1928 | 0.0030 | 0.1676 | 0.0024 | 0.1629 |
| $\widehat{F}_{C A L N E W O P T}$ | 0.0047 | 0.1895 | 0.0059 | 0.1928 | 0.0030 | 0.1676 | 0.0024 | 0.1629 |
|  | AVRB | AVRE | AVRB | AVRE | AVRB | AVRE | AVRB | AVRE |
|  | $n=40$ |  | $n=42$ |  | $n=45$ |  | $n=47$ |  |
| $\widehat{F}_{H T}$ | 0.0056 | 1 | 0.0032 | 1 | 0.0044 | 1 | 0.0052 | 1 |
| $\widehat{F}_{D}$ | 0.0160 | 0.4620 | 0.0102 | 0.4670 | 0.0089 | 0.4675 | 0.0108 | 0.4481 |
| $\widehat{F}_{R}$ | 0.0018 | 0.4241 | 0.0033 | 0.4380 | 0.0045 | 0.4389 | 0.0041 | 0.4160 |
| $\widehat{F}_{C D}$ | 0.1759 | 1.0417 | 0.1703 | 1.0957 | 0.1625 | 1.1100 | 0.1598 | 1.1054 |
| $\widehat{F}_{R K M}$ | 0.0076 | 0.4012 | 0.0023 | 0.4070 | 0.0011 | 0.4136 | 0.0017 | 0.4014 |
| $\widehat{F}_{C A L}$ | 0.0049 | 0.8386 | 0.0032 | 0.8394 | 0.0034 | 0.9319 | 0.0034 | 0.8700 |
| $\widehat{F}_{C A L 3}$ | 0.0008 | 0.3397 | 0.0024 | 0.3466 | 0.0033 | 0.3442 | 0.0022 | 0.3273 |
| $\widehat{F}_{C A L M A X}$ | 0.0043 | 0.1833 | 0.0028 | 0.1938 | 0.0031 | 0.1962 | 0.0036 | 0.1972 |
| $\widehat{F}_{C A L O P T}$ | 0.0022 | 0.1530 | 0.0013 | 0.1582 | 0.0011 | 0.1557 | 0.0019 | 0.1516 |
| $\widehat{F}_{C A L N E W O P T}$ | 0.0022 | 0.1530 | 0.0013 | 0.1582 | 0.0011 | 0.1557 | 0.0019 | 0.1516 |

Tables 3 and 4 provide the results obtained for the Simh population. From the results of Table 3 (AVrb and AVre) we can again observe that $\widehat{F}_{C A L O P T}(t)$ and $\widehat{F}_{C A L N E W O P T}(t)$ have the same behavior and they are the most efficient estimators for all sample sizes. Also, they also present a less bias for most of sample sizes. Additionally, we can highlight the bias and efficiency problems of $\widehat{F}_{C D}(t)$ because this estimator is biased when the relationship between $y$ and $x$ is not linear. As in the previous case, the dimension reduction analysis (Table 4 is essential to find out if $\widehat{F}_{C A L N E W O P T}(t)$ is a better alternative than $\widehat{F}_{C A L O P T}(t)$. In this case, from the results of Table 4 we can verify that again there is a slight reduction in the optimal vector used in $\widehat{F}_{\text {CALNEWOPT }}(t)$ for the smalls and medium quantiles. On the contrary, there is a moderate reduction for the higher quantiles where the dimension of the optimal vector used in $\widehat{F}_{C A L O P T}(t)$ has a value between 10 and 20 while the dimension for $\widehat{F}_{C A L N E W O P T}(t)$ remains between 2 and 5 for all sample sizes. Due to this reduction, the new estimator $\widehat{F}_{C A L N E W O P T}(t)$ provides a considerable benefit in execution time, especially in the higher quantiles but as in the previous case, this moderate reduction in the dimension of the auxiliary information used in the calibration procedure does not allow an improvement in the asymptotic efficiency. Probably, in this case we have a considerable cardinal for the set $Z_{t}$, although the set $F_{t}$ is not empty.

TABLE 2 Average dimension(MD), variance dimension (VD) and comparison of execution time (RT) of the estimators $\widehat{F}_{\text {CALOPT }}$ and $\widehat{F}_{\text {CALNEWOPT }}$. Population: DNase.

|  | $n=30$ |  |  |  |  | $n=32$ |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | MD |  | VD | RT | MD |  | VD |  | RT |
|  | $O P T$ | NEWOPT | OPT | NEWOPT |  | OPT | NEWOPT | OPT | NEWOPT |  |
| $t_{1}$ | 1 | 1 | 0 | 0 | 1 | 1 | 1 | 0 | 0 | 1 |
| $t_{2}$ | 1.929 | 1.698 | 0.257 | 0.459 | 0.875 | 1.920 | 1.742 | 0.271 | 0.438 | 0.556 |
| $t_{3}$ | 1.976 | 1 | 0.153 | 0 | 0.600 | 1.978 | 1 | 0.147 | 0 | 0.217 |
| $t_{4}$ | 2.781 | 1.738 | 0.419 | 0.440 | 0.846 | 2.807 | 1.761 | 0.405 | 0.427 | 0.467 |
| $t_{5}$ | 3.658 | 1.668 | 0.499 | 0.471 | 0.813 | 3.665 | 1.670 | 0.491 | 0.470 | 0.895 |
| $t_{6}$ | 3.951 | 1 | 0.225 | 0 | 0.238 | 3.960 | 1 | 0.196 | 0 | 0.412 |
| $t_{7}$ | 4.915 | 1.515 | 0.286 | 0.500 | 0.375 | 4.939 | 1.530 | 0.244 | 0.499 | 0.849 |
| $t_{8}$ | 5.883 | 1.739 | 0.343 | 0.439 | 0.579 | 5.897 | 1.763 | 0.307 | 0.425 | 0.261 |
| $t_{9}$ | 5.921 | 1 | 0.288 | 0 | 0.556 | 5.938 | 1 | 0.241 | 0 | 0.600 |
| $t_{10}$ | 6.723 | 1.733 | 0.476 | 0.443 | 0.571 | 6.765 | 1.756 | 0.452 | 0.430 | 0.909 |
| $t_{11}$ | 7.502 | 1.573 | 0.580 | 0.495 | 0.536 | 7.578 | 1.632 | 0.541 | 0.483 | 0.950 |
|  | $n=35$ |  |  |  |  | $n=37$ |  |  |  |  |
|  | MD |  | VD |  | RT | MD |  | VD |  | RT |
|  | $O P T$ | NEWOPT | OPT | NEWOPT |  | $O P T$ | NEWOPT | OPT | NEWOPT |  |
| $t_{1}$ | 1 | 1 | 0 | 0 | 1 | 1 | 1 | 0 | 0 | 0.187 |
| $t_{2}$ | 1.951 | 1.816 | 0.216 | 0.388 | 0.972 | 1.960 | 1.815 | 0.196 | 0.388 | 0.974 |
| $t_{3}$ | 1.989 | 1 | 0.104 | 0 | 0.379 | 1.990 | 1 | 0.100 | 0 | 0.250 |
| $t_{4}$ | 2.853 | 1.803 | 0.357 | 0.398 | 0.762 | 2.876 | 1.847 | 0.330 | 0.360 | 0.941 |
| $t_{5}$ | 3.714 | 1.714 | 0.465 | 0.452 | 0.650 | 3.749 | 1.742 | 0.438 | 0.438 | 0.615 |
| $t_{6}$ | 3.977 | 1 | 0.150 | 0 | 0.375 | 3.986 | 1 | 0.118 | 0 | 0.833 |
| $t_{7}$ | 4.962 | 1.608 | 0.196 | 0.488 | 0.524 | 4.971 | 1.572 | 0.168 | 0.495 | 0.773 |
| $t_{8}$ | 5.936 | 1.817 | 0.249 | 0.387 | 0.515 | 5.948 | 1.803 | 0.231 | 0.398 | 0.599 |
| $t_{9}$ | 5.969 | 1 | 0.173 | 0 | 0.125 | 5.975 | 1 | 0.162 | 0 | 0.471 |
| $t_{10}$ | 6.823 | 1.810 | 0.382 | 0.392 | 0.769 | 6.865 | 1.864 | 0.353 | 0.343 | 0.905 |
| $t_{11}$ | 7.650 | 1.671 | 0.494 | 0.470 | 0.667 | 7.654 | 1.677 | 0.497 | 0.468 | 0.714 |
|  | $n=40$ |  |  |  |  | $n=42$ |  |  |  |  |
|  | MD |  | VD |  | RT | MD |  | VD |  | RT |
|  | $O P T$ | NEWOPT | OPT | NEWOPT |  | OPT | NEWOPT | OPT | NEWOPT |  |
| $t_{1}$ | 1 | 1 | 0 | 0 | 1 | 1 | 0 | 0 | 1 |  |
| $t_{2}$ | 1.983 | 1.862 | 0.129 | 0.345 | 0.760 | 1.982 | 1.889 | 0.133 | 0.314 | 0.982 |
| $t_{3}$ | 1.998 | 1 | 0.0447 | 0 | 0.818 | 1.996 | 1 | 0.063 | 0 | 0.353 |
| $t_{4}$ | 2.898 | 1.871 | 0.303 | 0.335 | 0.706 | 2.920 | 1.895 | 0.271 | 0.307 | 0.753 |
| $t_{5}$ | 3.793 | 1.789 | 0.405 | 0.408 | 0.278 | 3.782 | 1.781 | 0.413 | 0.414 | 0.647 |
| $t_{6}$ | 3.996 | 1 | 0.063 | 0 | 0.346 | 3.993 | 1 | 0.083 | 0 | 0.500 |
| $t_{7}$ | 4.989 | 1.613 | 0.104 | 0.487 | 0.615 | 4.989 | 1.665 | 0.104 | 0.472 | 0.698 |
| $t_{8}$ | 5.974 | 1.869 | 0.159 | 0.338 | 0.587 | 5.979 | 1.880 | 0.143 | 0.325 | 0.440 |
| $t_{9}$ | 5.991 | 1 | 0.094 | 0 | 0 | 5.992 | 1 | 0.089 | 0.403 | 0.214 |
|  | 6.909 | 1.882 | 0.291 | 0.323 | 0.814 | 6.893 | 1.875 | 0.309 | 0.331 | 0.692 |
| $t_{11}$ | 7.731 | 1.735 | 0.446 | 0.441 | 0.773 | 7.750 | 1.756 | 0.433 | 0.430 | 0.400 |
|  | $n=45$ |  |  |  |  | $n=47$ |  |  |  |  |
|  | MD |  | VD |  | RT | MD |  | VD |  | RT |
|  | $O P T$ | NEWOPT | OPT | NEWOPT |  | $O P T$ | NEWOPT | $O P T$ | NEWOPT |  |
| $t_{1}$ | 1 | 1 | 0 | 0 | 1 | 1 | 1 | 0 | 0 | 1 |
| $t_{2}$ | 1.989 | 1.905 | 0.104 | 0.293 | 0.991 | 1.993 | 1.920 | 0.083 | 0.271 | 0.882 |
| $t_{3}$ | 1.998 | 1 | 0.045 | 0 | 0.467 | 1.999 | 1 | 0.032 | 0 | 0.600 |
| $t_{4}$ | 2.929 | 1.914 | 0.257 | 0.281 | 0.793 | 2.937 | 1.929 | 0.243 | 0.257 | 0.733 |
| $t_{5}$ | 3.834 | 1.829 | 0.372 | 0.377 | 0.882 | 3.863 | 1.863 | 0.347 | 0.344 | 0.805 |
| $t_{6}$ | 3.994 | 1 | 0.077 | 0 | 0.083 | 3.997 | 1 | 0.055 | 0 | 0.556 |
| $t_{7}$ | 4.990 | 1.698 | 0.100 | 0.459 | 0.786 | 4.995 | 1.705 | 0.071 | 0.456 | 0.529 |
| $t_{8}$ | 5.981 | 1.908 | 0.137 | 0.289 | 0.450 | 5.982 | 1.920 | 0.133 | 0.271 | 0.857 |
| $t_{9}$ | 5.993 | 1 | 0.083 | 0 | 0.136 | 5.995 | 1 | 0.071 | 0 | 0.368 |
| $t_{10}$ | 6.928 | 1.919 | 0.262 | 0.273 | 0.875 | 6.931 | 1.918 | 0.254 | 0.275 | 0.615 |
| $t_{11}$ | 7.790 | 1.794 | 0.415 | 0.405 | 0.782 | 7.782 | 1.784 | 0.416 | 0.412 | 0.643 |

TABLE 3 Average relative bias (AVRB) and average relative efficiency (AVRE) of the estimators compared. Population: Simh.

|  |  |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | AVRB | AVRE | AVRB | AVRE | AVRB | AVRE | AVRB | AVRE |
|  | $n=75$ | $n=100$ |  | $c$ | $n=125$ | $n=150$ |  |  |
| $\widehat{F}_{H T}$ | 0.0066 | 1 | 0.0018 | 1 | 0.0036 | 1 | 0.0034 | 1 |
| $\widehat{F}_{D}$ | 0.0075 | 0.5584 | 0.0016 | 0.5648 | 0.0027 | 0.5809 | 0.0041 | 0.5616 |
| $\widehat{F}_{R}$ | 0.0078 | 1.2980 | 0.0018 | 1.2368 | 0.0031 | 1.2648 | 0.0041 | 1.2889 |
| $\widehat{F}_{C D}$ | 0.3645 | 9.9689 | 0.3708 | 13.4606 | 0.3662 | 16.9590 | 0.3651 | 19.7880 |
| $\widehat{F}_{R K M}$ | 0.0080 | 0.3990 | 0.0055 | 0.3822 | 0.0038 | 0.3990 | 0.0044 | 0.3789 |
| $\widehat{F}_{C A L}$ | 0.0059 | 0.9369 | 0.0012 | 0.8205 | 0.0016 | 0.8794 | 0.0037 | 0.8845 |
| $\widehat{F}_{C A L 3}$ | 0.0028 | 0.3639 | 0.0015 | 0.3642 | 0.0008 | 0.3749 | 0.0037 | 0.3545 |
| $\widehat{F}_{C A L M A X}$ | 0.0012 | 0.2112 | 0.0008 | 0.1961 | 0.0016 | 0.1924 | 0.0009 | 0.1862 |
| $\widehat{F}_{C A L O P T}$ | 0.0030 | 0.1800 | 0.0012 | 0.1591 | 0.0006 | 0.1555 | 0.0010 | 0.1441 |
| $\widehat{F}_{C A L N E W O P T}$ | 0.0030 | 0.1800 | 0.0012 | 0.1591 | 0.0006 | 0.1555 | 0.0010 | 0.1441 |
|  | AVRB | AVRE | AVRB | AVRE | AVRB | AVRE | AVRB | AVRE |
|  | $n=175$ |  | $n=200$ |  | $n=250$ |  | $n=300$ |  |
| $\widehat{F}_{H T}$ | 0.0025 | 1 | 0.0037 | 1 | 0.0026 | 1 | 0.0015 | 1 |
| $\widehat{F}_{D}$ | 0.0025 | 0.5626 | 0.0032 | 0.5822 | 0.0013 | 0.5593 | 0.0011 | 0.5653 |
| $\widehat{F}_{R}$ | 0.0023 | 1.3001 | 0.0044 | 1.2704 | 0.0021 | 1.3077 | 0.0009 | 1.2659 |
| $\widehat{F}_{C D}$ | 0.3653 | 23.6861 | 0.3743 | 26.3930 | 0.3659 | 32.5335 | 0.3632 | 38.7995 |
| $\widehat{F}_{R K M}$ | 0.0040 | 0.3886 | 0.0024 | 0.3926 | 0.0020 | 0.3770 | 0.0011 | 0.3891 |
| $\widehat{F}_{C A L}$ | 0.0021 | 0.8935 | 0.0048 | 0.8732 | 0.0011 | 0.8603 | 0.0012 | 0.8778 |
| $\widehat{F}_{C A L 3}$ | 0.0020 | 0.3780 | 0.0022 | 0.3721 | 0.0011 | 0.3480 | 0.0007 | 0.3598 |
| $\widehat{F}_{C A L M A X}$ | 0.0007 | 0.1920 | 0.0008 | 0.1853 | 0.0008 | 0.1758 | 0.0006 | 0.1725 |
| $\widehat{F}_{C A L O P T}$ | 0.0003 | 0.1486 | 0.0007 | 0.1438 | 0.0008 | 0.1303 | 0.0005 | 0.1313 |
| $\widehat{F}_{C A L N E W O P T}$ | 0.0003 | 0.1486 | 0.0007 | 0.1438 | 0.0008 | 0.1303 | 0.0005 | 0.1313 |

For the Simser population, Tables 5 and 6 provide the results of the simulation study. In this case, Table 5 shows that $\widehat{F}_{C A L O P T}(t)$ and $\widehat{F}_{C A L N E W O P T}(t)$ do not present the same behavior and $\widehat{F}_{C A L N E W O P T}(t)$ is the one with the least bias and the best efficiency of all the estimators included in the simulation study and it produce a considerable improvement in efficiency with respect to $\widehat{F}_{C A L O P T}(t)$. For some of the estimators included in the simulation study, we can observe a worse efficiency than $\widehat{F}_{H T}(t)$, which may be caused by the absence of a linear relationship between $y$ and $x$. Regarding dimension reduction analysis and efficiency in execution time, Table 6 shows that for all quantiles, the optimal vector of $\widehat{F}_{C A L O P T}(t)$ increase its size when the sample size increases, especially in the larger quantiles, where we can observe very high dimensional optimal vectors. On the other hand, the dimension of the optimal vector for $\widehat{F}_{C A L N E W O P T}(t)$ remains stable for all sample sizes and it always shows a value below 7 and in most cases its value is between 3 and 5. It represents a quite considerable reduction of the dimension that causes a quite remarkable improvement in execution, especially in the high quantiles and according to previous studies ${ }^{23]}$ this remarkable reduction allows $\widehat{F}_{C A L N E W O P T}(t)$ to achieve an improvement in efficiency with respect to $\widehat{F}_{C A L O P T}(t)$. In this case, the cardinal of the set $Z_{t}$ is probably very high and it is similar to the cardinal of the set $A_{t}$, which implies that the set $F_{t}$ has few elements.

TABLE 4 Average dimension(MD), variance dimension (VD) and comparison of execution time (RT) of the estimators $\widehat{F}_{\text {CALOPT }}$ and $\widehat{F}_{\text {CALNEWOPT }}$. Population: Simh.

|  | $n=75$ |  |  |  |  | $n=100$ |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | MD |  |  | VD | RT | MD |  | VD |  | RT |
|  | OPT | NEWOPT | OPT | NEWOPT |  | OPT | NEWOPT | OPT | NEWOPT |  |
| $t_{1}$ | 2.355 | 0.543 | 1.395 | 0.489 | 0.557 | 2.471 | 0.527 | 1.489 | 0.500 | 0.780 |
| $t_{2}$ | 4.168 | 0.529 | 1.44 | 0.584 | 0.400 | 4.313 | 0.503 | 1.597 | 0.655 | 0.299 |
| $t_{3}$ | 5.18 | 0.560 | 1.55 | 0.639 | 0.710 | 5.341 | 0.522 | 1.739 | 0.676 | 0.541 |
| $t_{4}$ | 6.289 | 0.616 | 1.73 | 0.697 | 0.364 | 6.47 | 0.551 | 1.91 | 0.696 | 0.335 |
| $t_{5}$ | 8.552 | 0.657 | 2.277 | 0.739 | 0.519 | 8.769 | 0.533 | 2.471 | 0.675 | 0.343 |
| $t_{6}$ | 10.509 | 0.702 | 2.416 | 0.795 | 0.504 | 10.78 | 0.633 | 2.681 | 0.811 | 0.327 |
| $t_{7}$ | 12.514 | 0.863 | 2.831 | 0.998 | 0.367 | 12.833 | 0.769 | 3.162 | 0.956 | 0.388 |
| $t_{8}$ | 14.475 | 0.822 | 2.923 | 0.956 | 0.242 | 14.807 | 0.645 | 3.239 | 0.877 | 0.351 |
| $t_{9}$ | 15.664 | 0.928 | 3.1 | 1.0213 | 0.359 | 16.087 | 0.811 | 3.5 | 0.944 | 0.294 |
| $t_{10}$ | 16.71 | 0.978 | 3.375 | 1.143 | 0.249 | 17.163 | 0.923 | 3.827 | 1.140 | 0.273 |
| $t_{11}$ | 18.704 | 0.969 | 3.519 | 1.187 | 0.337 | 19.178 | 0.750 | 3.934 | 1.163 | 0.294 |
|  | $n=125$ |  |  |  |  | $n=150$ |  |  |  |  |
|  | MD |  |  | VD RT |  | MD |  |  | VD | RT |
|  | OPT | NEWOPT | OPT | NEWOPT |  | OPT | NEWOPT | OPT | NEWOPT |  |
| $t_{1}$ | 2.605 | 0.493 | 1.608 | 0.488 | 0.620 | 2.652 | 0.477 | 1.651 | 0.477 | 0.516 |
| $t_{2}$ | 4.392 | 0.493 | 1.695 | 0.671 | 0.610 | 4.471 | 0.501 | 1.841 | 0.692 | 0.291 |
| $t_{3}$ | 5.414 | 0.500 | 1.852 | 0.709 | 0.554 | 5.492 | 0.500 | 1.981 | 0.704 | 0.627 |
| $t_{4}$ | 6.58 | 0.5018 | 2.041 | 0.714 | 0.488 | 6.673 | 0.469 | 2.224 | 0.687 | 0.659 |
| $t_{5}$ | 8.883 | 0.456 | 2.64 | 0.628 | 0.339 | 8.961 | 0.414 | 2.775 | 0.573 | 0.409 |
| $t_{6}$ | 10.928 | 0.541 | 2.865 | 0.734 | 0.488 | 11.058 | 0.517 | 3.033 | 0.717 | 0.377 |
| $t_{7}$ | 12.988 | 0.658 | 3.415 | 0.892 | 0.343 | 13.092 | 0.621 | 3.578 | 0.876 | 0.342 |
| $t_{8}$ | 14.95 | 0.586 | 3.412 | 0.824 | 0.208 | 15.09 | 0.520 | 3.657 | 0.812 | 0.275 |
| $t_{9}$ | 16.277 | 0.753 | 3.698 | 0.936 | 0.308 | 16.394 | 0.708 | 3.862 | 0.896 | 0.304 |
| $t_{10}$ | 17.42 | 0.809 | 4.127 | 1.096 | 0.353 | 17.526 | 0.810 | 4.293 | 1.031 | 0.241 |
| $t_{11}$ | 19.438 | 0.641 | 4.333 | 1.089 | 0.262 | 19.532 | 0.584 | 4.564 | 1.087 | 0.278 |
|  | $n=175$ |  |  |  |  | $n=200$ |  |  |  |  |
|  | MD |  | VD |  | RT | MD |  |  | VD | RT |
|  | OPT | NEWOPT | $O P T$ | NEWOPT |  | OPT | NEWOPT | OPT | NEWOPT |  |
| $t_{1}$ | 2.719 | 0.450 | 1.719 | 0.450 | 0.602 | 2.763 | 0.426 | 1.763 | 0.426 | 0.603 |
| $t_{2}$ | 4.507 | 0.500 | 1.929 | 0.699 | 0.711 | 4.581 | 0.494 | 1.991 | 0.688 | 0.613 |
| $t_{3}$ | 5.551 | 0.498 | 2.056 | 0.712 | 0.450 | 5.588 | 0.492 | 2.167 | 0.701 | 0.277 |
| $t_{4}$ | 6.714 | 0.454 | 2.288 | 0.657 | 0.466 | 6.776 | 0.417 | 2.431 | 0.634 | 0.674 |
| $t_{5}$ | 9 | 0.364 | 2.858 | 0.508 | 0.372 | 9.022 | 0.360 | 2.901 | 0.485 | 0.366 |
| $t_{6}$ | 11.062 | 0.500 | 3.088 | 0.685 | 0.473 | 11.117 | 0.494 | 3.211 | 0.727 | 0.328 |
| $t_{7}$ | 13.178 | 0.617 | 3.74 | 0.848 | 0.440 | 13.214 | 0.631 | 3.846 | 0.835 | 0.355 |
| $t_{8}$ | 15.136 | 0.542 | 3.765 | 0.795 | 0.299 | 15.191 | 0.495 | 3.897 | 0.754 | 0.310 |
| $t_{9}$ | 16.474 | 0.694 | 4.003 | 0.859 | 0.303 | 16.576 | 0.667 | 4.209 | 0.830 | 0.326 |
| $t_{10}$ | 17.632 | 0.806 | 4.514 | 0.981 | 0.340 | 17.726 | 0.799 | 4.652 | 1.021 | 0.260 |
| $t_{11}$ | 19.617 | 0.524 | 4.747 | 1.021 | 0.268 | 19.652 | 0.523 | 4.891 | 1.009 | 0.227 |
|  | $n=250$ |  |  |  |  | $n=300$ |  |  |  |  |
|  | MD |  |  | VD | RT | MD |  |  | VD | RT |
|  | OPT | NEWOPT | OPT | NEWOPT |  | OPT | NEWOPT | OPT | NEWOPT |  |
| $t_{1}$ | 2.832 | 0.374 | 1.832 | 0.374 | 0.702 | 2.894 | 0.308 | 1.894 | 0.308 | 0.763 |
| $t_{2}$ | 4.659 | 0.474 | 2.135 | 0.691 | 0.522 | 4.714 | 0.452 | 2.29 | 0.668 | 0.519 |
| $t_{3}$ | 5.675 | 0.469 | 2.319 | 0.668 | 0.537 | 5.758 | 0.429 | 2.517 | 0.588 | 0.610 |
| $t_{4}$ | 6.858 | 0.349 | 2.596 | 0.563 | 0.480 | 6.93 | 0.255 | 2.712 | 0.479 | 0.336 |
| $t_{5}$ | 9.069 | 0.304 | 3.01 | 0.387 | 0.408 | 9.115 | 0.346 | 3.09 | 0.385 | 0.369 |
| $t_{6}$ | 11.178 | 0.465 | 3.319 | 0.685 | 0.453 | 11.242 | 0.467 | 3.392 | 0.702 | 0.297 |
| $t_{7}$ | 13.312 | 0.635 | 4.061 | 0.844 | 0.349 | 13.405 | 0.621 | 4.209 | 0.817 | 0.297 |
| $t_{8}$ | 15.268 | 0.490 | 4.056 | 0.702 | 0.272 | 15.357 | 0.506 | 4.211 | 0.699 | 0.325 |
| $t_{9}$ | 16.67 | 0.649 | 4.365 | 0.786 | 0.330 | 16.797 | 0.664 | 4.582 | 0.787 | 0.322 |
| $t_{10}$ | 17.897 | 0.766 | 4.914 | 0.980 | 0.310 | 18.041 | 0.735 | 5.179 | 0.904 | 0.304 |
| $t_{11}$ | 19.777 | 0.419 | 5.157 | 0.902 | 0.323 | 19.821 | 0.386 | 5.317 | 0.892 | 0.294 |

TABLE 5 Average relative bias (AVRB) and average relative efficiency (AVRE) of the estimators compared. Population: Simser.

|  | AVRB | AVRE | AVRB | AVRE | AVRB | AVRE | AVRB | AVRE |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $n=75$ |  | $n=100$ |  | $n=125$ |  | $n=150$ |  |
| $\widehat{F}_{H T}$ | 0.0042 | 1 | 0.0049 | 1 | 0.0042 | 1 | 0.0032 | 1 |
| $\widehat{F}_{D}$ | 0.0038 | 0.9963 | 0.0039 | 1.0008 | 0.0056 | 1.0022 | 0.0027 | 0.9999 |
| $\widehat{F}_{R}$ | 0.0832 | 3.0296 | 0.0687 | 3.0130 | 0.0659 | 2.8977 | 0.0564 | 3.0535 |
| $\widehat{F}_{C D}$ | 0.0305 | 0.8968 | 0.0215 | 0.9234 | 0.0196 | 0.9374 | 0.0117 | 0.9409 |
| $\widehat{F}_{R K M}$ | 0.0177 | 1.0419 | 0.0106 | 1.0246 | 0.0196 | 1.0530 | 0.0110 | 1.0252 |
| $\widehat{F}_{C A L}$ | 0.0271 | 1.5805 | 0.0199 | 1.5456 | 0.0274 | 1.5784 | 0.0201 | 1.5548 |
| $\widehat{F}_{C A L 3}$ | 0.0110 | 0.8375 | 0.0104 | 0.7839 | 0.0090 | 0.8138 | 0.0075 | 0.7860 |
| $\widehat{F}_{\text {CALMAX }}$ | 0.0174 | 0.6580 | 0.0146 | 0.6495 | 0.0229 | 0.6857 | 0.0177 | 0.6580 |
| $\widehat{F}_{\text {CALOPT }}$ | 0.0732 | 0.4007 | 0.0537 | 0.2894 | 0.0417 | 0.2307 | 0.0327 | 0.1799 |
| $\widehat{F}_{\text {CALNEWOPT }}$ | 0.0028 | 0.1516 | 0.0017 | 0.1143 | 0.0015 | 0.0985 | 0.0016 | 0.0787 |
|  | AVRB | AVRE | AVRB | AVRE | AVRB | AVRE | AVRB | AVRE |
|  | $n=175$ |  | $n=200$ |  | $n=250$ |  | $n=300$ |  |
| $\widehat{F}_{H T}$ | 0.0029 | 1 | 0.0026 | 1 | 0.0011 | 1 | 0.0023 | 1 |
| $\widehat{F}_{D}$ | 0.0037 | 1.0033 | 0.0022 | 0.9988 | 0.0013 | 1 | 0.0029 | 1.0016 |
| $\widehat{F}_{R}$ | 0.0547 | 3.0697 | 0.0461 | 3.0468 | 0.0449 | 3.0664 | 0.0418 | 3.0488 |
| $\widehat{F}_{C D}$ | 0.0094 | 0.9427 | 0.0139 | 0.9404 | 0.0127 | 0.9603 | 0.0121 | 0.9651 |
| $\widehat{F}_{R K M}$ | 0.0129 | 1.0307 | 0.0072 | 1.0161 | 0.0079 | 1.0221 | 0.0090 | 1.0152 |
| $\widehat{F}_{C A L}$ | 0.0215 | 1.5860 | 0.0143 | 1.4787 | 0.0160 | 1.5565 | 0.0169 | 1.5472 |
| $\widehat{F}_{C A L 3}$ | 0.0079 | 0.8379 | 0.0071 | 0.8280 | 0.0051 | 0.7523 | 0.0042 | 0.7902 |
| $\widehat{F}_{C A L M A X}$ | 0.0191 | 0.6812 | 0.0138 | 0.6426 | 0.0155 | 0.6591 | 0.0163 | 0.6747 |
| $\widehat{F}_{C A L O P T}$ | 0.0266 | 0.1529 | 0.0228 | 0.1248 | 0.0169 | 0.0900 | 0.0128 | 0.0733 |
| $\widehat{F}_{\text {CALNEWOPT }}$ | 0.0010 | 0.0714 | 0.0007 | 0.0577 | 0.0007 | 0.0476 | 0.0004 | 0.0424 |

## 6 | DISCUSSION AND CONCLUSIONS

In recent years, the calibration technique has attracted significant attention in survey sampling research and survey applications. The calibration method allows obtaining more reliable estimates for a finite population by incorporating auxiliary information available in the population.

In this article, we investigate whether the optimal estimator in the proposal ${ }^{33}$ (that can be applied direclty in the estimation of qunatile and poverty measures ${ }^{(21)}$ ) based on the calibration method for estimating the distribution function can be improved by reducing the dimension of the optimal vector used in the calibration process. Working with a reduced number of variables may reduce numerical problems related to optimization procedures and also limit the presence of negative, very large and unstable calibration weights. To do this, we have theoretically established the conditions under which a reduction in the dimension of the optimal vector is possible and through an extensive simulation study we have verified how the new estimator $\widehat{F}_{\text {CALNEWOPT }}(t)$ can avoid the problems associated with a high-dimensional auxiliary data and allows to improve the execution time maintaining (DNase and Simh) or even improving the efficiency ${ }^{[23]}$ (Simser). Therefore, the new proposal is a more reliable option when carrying out real analyzes where large population sizes can lead to high-dimensional optimal vectors for $\widehat{F}_{C A L O P T}(t)$, while $\widehat{F}_{C A L N E W O P T}(t)$ can lead to a considerable reduction in this optimal dimension.

Further research is needed regarding the dimension reduction on calibration for the distribution function as our study presents certain limitations. Our paper is restricted to a simple random sampling design. The determination of the optimal vector for calibration (and its dimension) can be extended relatively easily to the case of self-weighted samples (for example, stratified samples

TABLE 6 Average dimension(MD), variance dimension (VD) and comparison of execution time (RT) of the estimators $\widehat{F}_{\text {CALOPT }}$ and $\widehat{F}_{\text {CALNEWOPT }}$. Population: Simser.

|  | $n=75$ |  |  |  |  | $n=100$ |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | MD |  |  | VD | RT | MD |  | VD |  | RT |
|  | OPT | NEWOPT | OPT | NEWOPT |  | OPT | NEWOPT | OPT | NEWOPT |  |
| $t_{1}$ | 7.807 | 2.355 | 3.036 | 0.277 | 1.152 | 9.846 | 2.589 | 3.084 | 0.376 | 0.779 |
| $t_{2}$ | 14.767 | 3.192 | 3.013 | 0.122 | 0.431 | 18.586 | 3.430 | 3.029 | 0.200 | 0.516 |
| $t_{3}$ | 18.115 | 3.395 | 3.006 | 0.077 | 0.449 | 23.052 | 3.825 | 3.022 | 0.166 | 0.443 |
| $t_{4}$ | 21.519 | 3.554 | 3.007 | 0.083 | 0.378 | 27.492 | 3.990 | 3.005 | 0.071 | 0.427 |
| $t_{5}$ | 28.379 | 3.778 | 3.002 | 0.045 | 0.418 | 36.322 | 4.225 | 3.002 | 0.045 | 0.39 |
| $t_{6}$ | 35.189 | 4.032 | 3.005 | 0.071 | 0.4 | 45.339 | 4.389 | 3.013 | 0.113 | 0.318 |
| $t_{7}$ | 42.003 | 4.073 | 3.015 | 0.122 | 0.296 | 54.139 | 4.445 | 3.005 | 0.071 | 0.283 |
| $t_{8}$ | 48.955 | 3.807 | 3 | 0 | 0.305 | 63.065 | 4.240 | 3.006 | 0.077 | 0.294 |
| $t_{9}$ | 52.367 | 3.655 | 3.003 | 0.055 | 0.338 | 67.537 | 4.076 | 3.009 | 0.094 | 0.237 |
| $t_{10}$ | 55.801 | 3.565 | 3.004 | 0.063 | 0.246 | 71.983 | 3.950 | 3.002 | 0.089 | 0.248 |
| $t_{11}$ | 62.753 | 3.023 | 2.948 | 0.363 | 0.228 | 80.714 | 3.531 | 2.982 | 0.289 | 0.268 |
|  | $n=125$ |  |  |  |  | $n=150$ |  |  |  |  |
|  | MD |  |  | VD RT |  | MD |  |  | VD | RT |
|  | OPT | NEWOPT | OPT | NEWOPT |  | OPT | NEWOPT | OPT | NEWOPT |  |
| $t_{1}$ | 11.934 | 3.145 | 2.664 | 0.471 | 0.946 | 13.766 | 2.767 | 3.141 | 0.455 | 0.762 |
| $t_{2}$ | 22.63 | 3.046 | 3.697 | 0.241 | 0.331 | 26.001 | 3.740 | 3.066 | 0.303 | 0.453 |
| $t_{3}$ | 27.879 | 3.023 | 3.967 | 0.169 | 0.42 | 32.224 | 4.073 | 3.027 | 0.185 | 0.337 |
| $t_{4}$ | 33.406 | 3.014 | 4.195 | 0.118 | 0.348 | 38.628 | 4.411 | 3.013 | 0.113 | 0.379 |
| $t_{5}$ | 44.154 | 3.011 | 4.626 | 0.122 | 0.325 | 51.218 | 4.853 | 3.008 | 0.090 | 0.287 |
| $t_{6}$ | 54.934 | 3.011 | 4.893 | 0.104 | 0.283 | 63.845 | 5.074 | 3.026 | 0.159 | 0.259 |
| $t_{7}$ | 65.748 | 3.01 | 4.875 | 0.010 | 0.253 | 76.315 | 5.219 | 3.019 | 0.137 | 0.243 |
| $t_{8}$ | 76.511 | 3.01 | 4.815 | 0.010 | 0.191 | 88.878 | 5.144 | 3.006 | 0.077 | 0.19 |
| $t_{9}$ | 81.862 | 3.01 | 4.695 | 0.010 | 0.255 | 95.292 | 5.014 | 3.013 | 0.113 | 0.212 |
| $t_{10}$ | 87.184 | 3.013 | 4.555 | 0.113 | 0.192 | 101.589 | 4.937 | 3.02 | 0.140 | 0.218 |
| $t_{11}$ | 98.008 | 3 | 4.022 | 0.219 | 0.208 | 114.415 | 4.415 | 3.027 | 0.180 | 0.195 |
|  | $n=175$ |  |  |  |  | $n=200$ |  |  |  |  |
|  | MD |  |  | VD | RT | MD |  |  | VD | RT |
|  | OPT | NEWOPT | OPT | NEWOPT |  | OPT | NEWOPT | OPT | NEWOPT |  |
| $t_{1}$ | 15.496 | 3.036 | 2.927 | 0.277 | 0.855 | 16.921 | 2.936 | 3.258 | 0.631 | 0.663 |
| $t_{2}$ | 29.729 | 3.013 | 3.853 | 0.122 | 0.434 | 32.618 | 3.922 | 3.089 | 0.351 | 0.367 |
| $t_{3}$ | 36.67 | 3.006 | 4.144 | 0.077 | 0.438 | 40.25 | 4.137 | 3.058 | 0.277 | 0.26 |
| $t_{4}$ | 43.817 | 3.007 | 4.441 | 0.083 | 0.335 | 48.197 | 4.458 | 3.037 | 0.189 | 0.276 |
| $t_{5}$ | 58.001 | 3.002 | 4.757 | 0.045 | 0.268 | 63.872 | 4.957 | 3.012 | 0.109 | 0.26 |
| $t_{6}$ | 72.284 | 3.005 | 5.065 | 0.071 | 0.227 | 79.743 | 5.293 | 3.033 | 0.179 | 0.234 |
| $t_{7}$ | 86.505 | 3.015 | 5.062 | 0.122 | 0.25 | 95.683 | 5.487 | 3.061 | 0.239 | 0.209 |
| $t_{8}$ | 100.946 | 3 | 5.083 | 0 | 0.219 | 111.761 | 5.488 | 3.011 | 0.104 | 0.191 |
| $t_{9}$ | 108.154 | 3.003 | 4.976 | 0.055 | 0.194 | 119.62 | 5.366 | 3.022 | 0.147 | 0.195 |
| $t_{10}$ | 115.308 | 3.004 | 4.883 | 0.063 | 0.18 | 127.542 | 5.398 | 3.031 | 0.173 | 0.171 |
| $t_{11}$ | 129.69 | 2.948 | 4.693 | 0.363 | 0.175 | 143.597 | 5.105 | 3.056 | 0.230 | 0.149 |
|  | $n=250$ |  |  |  |  | $n=300$ |  |  |  |  |
|  | MD |  |  | VD | RT | MD |  |  | VD RT <br> NEWOPT  |  |
|  | OPT | NEWOPT | OPT | NEWOPT |  | OPT | NEWOPT OPT |  |  |  |  |
| $t_{1}$ | 20.073 | 3.438 | 3.111 | 0.787 | 0.552 | 22.612 | 3.533 | 3.117 | 0.860 | 0.646 |
| $t_{2}$ | 38.419 | 3.173 | 4.275 | 0.491 | 0.294 | 43.526 | 3.193 | 4.048 | 0.484 | 0.381 |
| $t_{3}$ | 47.492 | 3.091 | 4.653 | 0.342 | 0.29 | 53.736 | 3.131 | 4.393 | 0.387 | 0.323 |
| $t_{4}$ | 57.048 | 3.054 | 5.007 | 0.226 | 0.288 | 64.659 | 3.072 | 4.745 | 0.259 | 0.301 |
| $t_{5}$ | 75.652 | 3.015 | 5.438 | 0.122 | 0.25 | 85.713 | 3.026 | 5.212 | 0.177 | 0.206 |
| $t_{6}$ | 94.572 | 3.058 | 5.786 | 0.234 | 0.204 | 107.232 | 3.081 | 5.639 | 0.273 | 0.191 |
| $t_{7}$ | 113.331 | 3.058 | 5.855 | 0.234 | 0.186 | 128.379 | 3.094 | 5.858 | 0.292 | 0.177 |
| $t_{8}$ | 132.281 | 3.024 | 5.856 | 0.153 | 0.174 | 149.69 | 3.035 | 5.987 | 0.184 | 0.15 |
| $t_{9}$ | 141.666 | 3.035 | 5.778 | 0.184 | 0.156 | 160.366 | 3.036 | 6.071 | 0.186 | 0.136 |
| $t_{10}$ | 150.953 | 3.043 | 5.768 | 0.203 | 0.14 | 171.005 | 3.067 | 6.076 | 0.250 | 0.135 |
| $t_{11}$ | 169.593 | 3.083 | 5.433 | 0.276 | 0.124 | 192.424 | 3.096 | 5.887 | 0.295 | 0.132 |

with proportional allocation). However, the case of sampling with unequal probabilities is more complex and the methodology to be used is not the same. In future research we try to extend the results of this paper from SRSWOR to complex sampling designs.

Another limitation of our work is that the estimator considered is based on a pseudo-variable $g_{k}$ that assumes a linear relationship between variable $y$ and the covariates. The selection of the optimal auxiliary vector for the estimators based on a non-linear model should be considered in future studies.

## Financial disclosure

This study was partially supported by Ministerio de Educación y Ciencia (PID2019-106861RB-I00, Spain), IMAG-Maria de Maeztu CEX2020-001105-M/AEI/10.13039/501100011033 and FEDER/Junta de Andalucía-Consejería de Transformación Económica, Industria, Conocimiento y Universidades (FQM170-UGR20).

## Conflict of interest

The authors declare no potential conflict of interests.

How to cite this article: S. Martínez, M. Rueda, M.D. Illescas (2021), Reduction of optimal calibration dimension with a new optimal auxiliary vector for calibrated estimators of the distribution function, Math. Method. Appl. Sci., 2021;00:1-6.

## APPENDIX

## A SUPPLEMENTARY CASES FOR SECTION 3

## A. 1 Dimension reduction of the optimal auxiliary vector for $t=y_{\text {max }}$

If we consider $t=y_{\text {max }}$ where

$$
y_{\max }=\max _{k \in U} y_{k}
$$

the set $A_{t}$ is given by

$$
\begin{equation*}
A_{t}=\left\{g_{k}: k \in U ; y_{k} \leq t\right\}=\left\{a_{1}, a_{2}, \ldots, a_{M}\right\}=A_{M} \tag{A1}
\end{equation*}
$$

with $a_{1}<a_{2}<\cdots<a_{M}$ and $M$ denotes the total number of different values that the pseudo variable $g$ can take in the population $U$. As a consequence, $B_{t}=\emptyset$, the optimal dimension $P=M$ is the highest value and the optimal vector is given by:

$$
\mathbf{t}_{\mathbf{O P}}(t)=\left(a_{1}, a_{2}, \ldots, a_{M}\right)
$$

Our purpose is to analyze the possibility of obtaining the minimum value of 14 by means of a lower-dimensional auxiliary vector. Firstly, if we consider the value $t=y_{\max }$, we have:

$$
K_{t}\left(a_{j}\right)=\sum_{k \in U} \Delta\left(a_{j}-g_{k}\right) \Delta\left(y_{\max }-y_{k}\right)=N \cdot F_{g}\left(a_{j}\right) \quad j=1, \ldots, M
$$

and where we set $a_{0}$ so that $F_{g}\left(a_{0}\right)=0$ and $K_{t}\left(a_{0}\right)$.
The value of $Q_{t}(\boldsymbol{\gamma})$ at $\boldsymbol{\gamma}=\mathbf{t}_{\mathbf{O P}}(t)$ is given by:

$$
\begin{gathered}
Q_{t}\left(\mathbf{t}_{\mathbf{O P}}(t)\right)=Q_{t}\left(a_{1}, a_{2} \ldots, a_{M}\right)=2 N F_{y}\left(y_{\max }\right) \cdot K_{t}\left(a_{M}\right)-\sum_{j=1}^{M} \frac{\left(K_{t}\left(a_{j}\right)-K_{t}\left(a_{j-1}\right)\right)^{2}}{\left(F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)\right)}-\left(K_{t}\left(a_{M}\right)\right)^{2}= \\
=2 N F_{y}\left(y_{\max }\right) \cdot N \cdot F_{g}\left(a_{M}\right)-\sum_{j=1}^{M} \frac{\left(F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)\right)^{2}}{\left(F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)\right)}-\left(N \cdot F_{g}\left(a_{M}\right)\right)^{2}= \\
=2 N F_{y}\left(y_{\max }\right) \cdot N \cdot F_{g}\left(a_{M}\right)-\sum_{j=1}^{M} N^{2} \cdot\left(F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)\right)-\left(N \cdot F_{g}\left(a_{M}\right)\right)^{2}=2 N^{2} F_{y}\left(y_{\max }\right) \cdot F_{g}\left(a_{M}\right)-\left(N \cdot F_{g}\left(a_{M}\right)\right)^{2} .
\end{gathered}
$$

Since $F_{y}\left(y_{\max }\right)=F_{g}\left(a_{M}\right)=1$, it is clear that $Q_{t}\left(\mathbf{t}_{\mathbf{O P}}(t)\right)=0$ and consequently the minimum value of $Q_{t}(\boldsymbol{\gamma})$ for $y_{\max }$ is equal to 0 .
On the other hand, if we consider $\gamma=\left(a_{M}\right)$ then

$$
\begin{gathered}
Q_{t}\left(a_{M}\right)=2 N F_{y}\left(y_{\max }\right) \cdot K_{t}\left(a_{M}\right)-\frac{\left(K_{t}\left(a_{M}\right)\right)^{2}}{\left(F_{g}\left(a_{j}\right)\right)}-\left(K_{t}\left(a_{M}\right)\right)^{2}= \\
Q_{t}\left(a_{M}\right)=2 N^{2} F_{y}\left(y_{\max }\right) \cdot F_{g}\left(a_{M}\right)-2 N^{2}\left(F_{g}\left(a_{M}\right)\right)^{2}=0
\end{gathered}
$$

Thus, with the auxiliary vector $\boldsymbol{\gamma}=\left(a_{M}\right)$ the minimum value of $Q_{t}(\boldsymbol{\gamma})$ is reached and the optimal dimension can be reducted from $M$ to 1 . With the auxiliary vector $\gamma=\left(a_{M}\right)$, the resulting calibration constraint is given by:

$$
\begin{equation*}
1=F_{g}\left(a_{M}\right)=\frac{1}{N} \sum_{k \in s} \omega_{k} \Delta\left(a_{M}-g_{k}\right)=\frac{1}{N} \sum_{k \in s} \omega_{k} \tag{A2}
\end{equation*}
$$

Under simple random sampling without replacement, the minimization of (5) subject to the condition (A2) results in $d_{k}=\omega_{k}$ since the basic weights $d_{k}$ associated with the simple random sampling without replacement satisfy the condition A2). Therefore, the minimum value of $Q_{t}(\boldsymbol{\gamma})$ for $y_{\max }$ is equal to 0 .

## A. 2 Dimension reduction of the optimal auxiliary vector when $D_{t}=\emptyset ; \mathbf{Z}_{t}=\emptyset$ and $F_{t}=A_{t}=A_{M}$

If we consider the case where $D_{t}=\emptyset ; Z_{t}=\emptyset$ and $F_{t}=A_{t}=A_{M}$ there is not reduction in the optimal auxiliary vector $\mathbf{t}_{\mathbf{O P}}(t)$. To see it, it is clear that $q_{i}^{t} \neq 0 \forall a_{i} \in A_{M}$ and consequently:

$$
K_{t}\left(a_{i}\right)=\sum_{\mathrm{i} n U} \Delta\left(a_{i}-g_{k}\right) \Delta\left(t-y_{k}\right)=N \cdot F_{g}\left(a_{i}\right)-\sum_{j=1}^{i} q_{j}^{t} \text { for } i=1,2, \ldots, M
$$

Specifically, for $i=M$ we have:

$$
K_{t}\left(a_{M}\right)=N \cdot F_{g}\left(a_{M}\right)-\sum_{j=1}^{i} q_{j}^{t}=N F_{y}(t)
$$

The minimum value of $Q_{t}(\gamma)$ is reached at $\gamma=\mathbf{t}_{\mathbf{O P}}(t)$ and is given by:

$$
Q_{t}\left(\mathbf{t}_{\mathbf{O P}}(t)\right)=2 N F_{y}(t) \cdot K_{t}\left(a_{M}\right)-\sum_{j=1}^{M} \frac{\left(K_{t}\left(a_{j}\right)-K_{t}\left(a_{j-1}\right)\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}-\left(K_{t}\left(a_{M}\right)\right)^{2}
$$

Since $\left(K_{t}\left(a_{j}\right)-K_{t}\left(a_{j-1}\right)=N \cdot F_{g}\left(a_{j}\right)-N \cdot F_{g}\left(a_{j-1}\right)-q_{j}^{t}, Q_{t}\left(\mathbf{t}_{\mathbf{O P}}(t)\right)\right.$ takes the following expression:

$$
\begin{gather*}
Q_{t}\left(\mathbf{t}_{\mathbf{O P}}(t)\right)=\left(N F_{y}(t)\right)^{2}-\sum_{j=1}^{M} \frac{\left(N \cdot F_{g}\left(a_{j}\right)-N \cdot F_{g}\left(a_{j-1}\right)-q_{j}^{t}\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}= \\
\left(N F_{y}(t)\right)^{2}-N^{2} \cdot \sum_{j=1}^{M}\left(F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)\right)+2 N \cdot \sum_{j=1}^{M} q_{j}^{t}-\sum_{j=1}^{M} \frac{\left(q_{j}^{t}\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}= \\
\left(N F_{y}(t)\right)^{2}-N^{2} \cdot F_{g}\left(a_{M}\right)+2 N \cdot \sum_{j=1}^{M} q_{j}^{t}-\sum_{j=1}^{M} \frac{\left(q_{j}^{t}\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}=\left(N F_{y}(t)\right)^{2}-N^{2}+2 N \cdot \sum_{j=1}^{M} q_{j}^{t}-\sum_{j=1}^{M} \frac{\left(q_{j}^{t}\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)} . \tag{A3}
\end{gather*}
$$

If we consider an auxiliary vector where we delete some value $a_{i} \neq a_{M}$, i.e $\gamma=\left(a_{1}, \ldots a_{i-1}, a_{i+1}, \ldots, a_{M}\right)$, we can obtain in a similar way that

$$
\begin{gathered}
Q_{t}(\gamma)=2 N F_{y}(t) \cdot K_{t}\left(a_{M}\right)-\sum_{j=1}^{i-1} \frac{\left(K_{t}\left(a_{j}\right)-K_{t}\left(a_{j-1}\right)\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}-\sum_{j=i+1}^{M} \frac{\left(K_{t}\left(a_{j}\right)-K_{t}\left(a_{j-1}\right)\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}-\frac{\left(K_{t}\left(a_{i+1}\right)-K_{t}\left(a_{i-1}\right)\right)^{2}}{F_{g}\left(a_{i+1}\right)-F_{g}\left(a_{i-1}\right)}-\left(K_{t}\left(a_{M}\right)\right)^{2}= \\
=\left(N F_{y}(t)\right)^{2}-N^{2}+2 N \cdot \sum_{\substack{j=1 \\
j \neq i, i+1}}^{M} q_{j}^{t}+2 N\left(q_{i}^{t}+q_{i+1}^{t}\right)-\sum_{\substack{j=1 \\
j \neq i, i+1}}^{M} \frac{\left(q_{j}^{t}\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}-\frac{\left(q_{i}^{t}+q_{i+1}^{t}\right)^{2}}{F_{g}\left(a_{i+1}\right)-F_{g}\left(a_{i-1}\right)} .
\end{gathered}
$$

As a consequence, $Q_{t}\left(\mathbf{t}_{\mathbf{O P}}(t)\right)-Q_{t}(\gamma)$ is given by

$$
\begin{gathered}
Q_{t}\left(\mathbf{t}_{\mathbf{O P}}(t)\right)-Q_{t}(\gamma)=-\frac{\left(q_{i}^{t}\right)^{2}}{F_{g}\left(a_{i}\right)-F_{g}\left(a_{i-1}\right)}-\frac{\left(q_{i+1}^{t}\right)^{2}}{F_{g}\left(a_{i+1}\right)-F_{g}\left(a_{i}\right)}+\frac{\left(q_{i}^{t}+q_{i+1}^{t}\right)^{2}}{F_{g}\left(a_{i+1}\right)-F_{g}\left(a_{i-1}\right)}= \\
=-\frac{\left(q_{i}^{t}\right)^{2}\left(F_{g}\left(a_{i+1}\right)-F_{g}\left(a_{i}\right)\right)}{\left(F_{g}\left(a_{i}\right)-F_{g}\left(a_{i-1}\right)\right)\left(\left(F_{g}\left(a_{i+1}\right)-F_{g}\left(a_{i-1}\right)\right)\right.}-\frac{\left(q_{i+1}^{t}\right)^{2}\left(F_{g}\left(a_{i}\right)-F_{g}\left(a_{i-1}\right)\right)}{\left(F_{g}\left(a_{i+1}\right)-F_{g}\left(a_{i-1}\right)\right)\left(\left(F_{g}\left(a_{i+1}\right)-F_{g}\left(a_{i}\right)\right)\right.}+\frac{2 q_{i}^{t} \cdot q_{i+1}^{t}}{F_{g}\left(a_{i+1}\right)-F_{g}\left(a_{i-1}\right)}= \\
=\Gamma \cdot\left[-\left(q_{i}^{t}\right)^{2}\left(F_{g}\left(a_{i+1}\right)-F_{g}\left(a_{i}\right)\right)^{2}-\left(q_{i+1}^{t}\right)^{2}\left(F_{g}\left(a_{i}\right)-F_{g}\left(a_{i-1}\right)\right)^{2}+2 q_{i}^{t} \cdot q_{i+1}^{t}\left(F_{g}\left(a_{i+1}\right)-F_{g}\left(a_{i}\right)\right)\left(F_{g}\left(a_{i}\right)-F_{g}\left(a_{i-1}\right)\right)\right]<0
\end{gathered}
$$

with

$$
\Gamma_{=} \frac{1}{\left(F_{g}\left(a_{i}\right)-F_{g}\left(a_{i-1}\right)\right)\left(F_{g}\left(a_{i+1}\right)-F_{g}\left(a_{i}\right)\right)\left(F_{g}\left(a_{i+1}\right)-F_{g}\left(a_{i-1}\right)\right)} .
$$

Consequently, when deleting some $a_{i}, Q_{t}\left(\mathbf{t}_{\mathbf{O P}}(t)\right)<Q_{t}(\gamma)$.
If we delete the value $a_{M}$, i.e, we consider the auxiliary vector $\gamma=\left(a_{1}, \ldots, a_{M-1}\right), Q_{t}(\gamma)$ takes the following expression:

$$
Q_{t}(\gamma)=2 N F_{y}(t) \cdot K_{t}\left(a_{M-1}\right)-\sum_{j=1}^{M-1} \frac{\left(K_{t}\left(a_{j}-K_{t}\left(a_{j-1}\right)\right)\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}-\left(K_{t}\left(a_{M-1}\right)\right)^{2}
$$

On the other hand

$$
Q_{t}\left(\mathbf{t}_{\mathbf{O}}(t)\right)=\left(N F_{y}(t)\right)^{2}-\sum_{j=1}^{M} \frac{\left(K_{t}\left(a_{j}-K_{t}\left(a_{j-1}\right)\right)\right)^{2}}{F_{g}\left(a_{j}\right)-F_{g}\left(a_{j-1}\right)}
$$

and it easy to see that

$$
Q_{t}\left(\mathbf{t}_{\mathbf{O P}}(t)\right)-Q_{t}(\gamma)=\left(N F_{y}(t)-K_{t}\left(a_{M-1}\right)\right)^{2}-\frac{\left(N F_{y}(t)-K_{t}\left(a_{M-1}\right)\right)^{2}}{F_{g}\left(a_{M}\right)-F_{g}\left(a_{M-1}\right)}<0 .
$$

Therefore, when we delete $a_{M} ; Q_{t}\left(\mathbf{t}_{\mathbf{O P}}(t)\right)<Q_{t}(\gamma)$.
Thus, if $Z_{t}=\emptyset$ and $F_{t}=A_{t}$ there is not a reduction in the auxiliary vector to reach the minimum of $Q_{t}(\gamma)$.

## A. 3 Dimension reduction of the optimal auxiliary vector for $p_{i} ; i \in\left\{2, \ldots, l_{t}\right\}$ when $\mathbf{D}_{\mathbf{t}} \neq \emptyset ; \mathbf{D}_{\mathbf{t}}=\mathbf{A}_{\mathbf{M}}$ and $B_{t} \neq \emptyset$

Under the assumptions $\mathbf{D}_{\mathbf{t}} \neq \emptyset ; \mathbf{D}_{\mathbf{t}}=\mathbf{A}_{\mathbf{M}}$ and $\mathbf{B}_{\mathbf{t}} \neq \emptyset$, if we consider $p_{i}$ with $i=2, \ldots, l_{t}$, it is clear that $p_{i}>p_{(i-1)}$ and $f_{p_{i}}^{t}>f_{p_{(-1)}}^{t}$. Moreover, because the value $b_{f_{p_{i}}}^{t}$ exists, this implies that $f_{p_{i}}^{t}>f_{p_{(i-1)}}^{t}+1$.

If we suppose that $p_{i}=p_{(i-1)}+1$, it is clear that $f_{p_{i}}^{t}=f_{\left(p_{(i-1)}+1\right)}^{t}$ and due to the value $b_{f_{p_{i}}}^{t}$ exists, the set

$$
U_{p_{i(-1)}+1}=U_{p_{i}}=\left\{l \in U: a_{f_{p_{(i-1)}}}<g_{l}<a_{f_{p_{(i-1)+1}^{\prime}}^{t}}\right\}=\left\{l \in U: a_{f_{p_{(i-1)}}}<g_{l}<a_{f_{t_{i}}}\right\} \neq \emptyset
$$

As a consequence, we have:

$$
\left\{a_{f_{\left(p_{(i-1)}+1\right)}}, \ldots, a_{\left(f_{p_{i}}^{\prime}-1\right)}\right\} \subseteq D_{t}
$$

and $b_{f_{p_{i}}}^{t}=a_{\left(f_{p_{i}}-1\right)}$ and there is not a possible reduction in the dimension.
On the contrary, if we suppose that $p_{i}>p_{(i-1)}+1$, then $f_{p_{i}}^{t}>f_{\left(p_{(i-1)}+1\right)}^{t}$ and there is a integer $z \geq 1$ such that $p_{i}=p_{(i-1)}+1+z$. For all $j=1, \ldots, z$, the value $b_{f_{\left.p_{(i-1)}\right)^{\prime}}}$ does not exist and the set $U_{p_{(i-1)}+j}^{(i-1)+1}=\emptyset$. As in the previous case (case $p_{1}$ ), we have:

$$
a_{f_{\left.f_{(i-1)}+1\right)}^{t}}=a_{\left(f_{p_{(i-1)}}^{t}+j\right)}, \quad j=1, \ldots, z .
$$

Thus, if $p_{i}>p_{(i-1)}+1$, we have:

$$
\left.\left.\left\{a_{\left(f_{p_{(-1)}}^{t}\right.}+1\right), \ldots, a_{\left(f_{p_{(i-1)}}^{t}\right.}+p_{i}-p_{(i-1)}-1\right)\right\} \subseteq A_{t} .
$$

Then, if we define the following sets:

$$
\begin{aligned}
A_{p_{i}}= & \left\{a_{\left(f_{p_{(i-1)}^{\prime}}\right.}+1, \ldots, a_{\left(f_{\left(p_{(-1)}^{\prime}\right)}+p_{i}-p_{(i-1)}-1\right)}\right\} \\
& Z_{p_{i}}=\left\{a_{i} \in A_{p_{i}}: q_{i}^{t}=0\right\}
\end{aligned}
$$

and

$$
F_{p_{i}}=\left\{a_{i} \in A_{p_{i}}: 0<q_{i}^{t}<r_{i}\right\}
$$

we can proof in a similar way to the previous case (case $p_{1}$ ) that if $Z_{p_{i}}=A_{p_{i}}$ or $Z_{p_{i}} \neq A_{p_{i}}$ but $F_{p_{i}} \neq A_{p_{i}}$ there is a possible reduction in the dimension of the auxiliary vector $\mathbf{t}_{\mathbf{O P}}(t)$. If $F_{p_{i}}=A_{p_{i}}$ there is no possible dimension reduction.

Finally, if we suppose that $p_{l_{t}}=M_{t}$ then the value $b_{f_{M_{t}}}$ exists and analogously to previous cases, there is no reduction between the points $b_{f_{M_{t}}^{t}}$ and $a_{f_{M_{t}}^{t}}$.

On the other hand, if we suppose that $p_{l_{t}}<M_{t}$ then for $h=p_{l_{t}}+1, p_{l_{t}}+2, \ldots, M_{t}$ the corresponding value $b_{f_{h}^{t}}$ does not exist and therefore the sets $U_{h}=\emptyset$. As a consequence, we have:

$$
a_{f_{\left(p_{t}+1\right)}^{t}}=a_{\left(f_{p_{l_{t}}}^{t}+1\right)}, \ldots, a_{f_{M_{t}}^{t}}=a_{\left(f_{p_{l_{t}}}^{t}+M_{t}-p_{l_{t}}\right)}
$$

If we denote by

$$
\begin{gathered}
A_{M_{t}}=\left\{a_{\left(f_{p_{t}}^{t}+1\right)}, \ldots, a_{\left(f_{p_{t}}+M_{t}-p_{t_{t}}\right)}\right\} \subseteq A_{t} \\
Z_{M_{t}}=\left\{a_{i} \in A_{M_{t}}: q_{i}^{t}=0\right\} \\
F_{M_{t}}=\left\{a_{i} \in A_{M_{t}}: 0<q_{i}^{t}<r_{i}\right\}
\end{gathered}
$$

then, we can reduce the dimension of the optimal auxiliary vector $\mathbf{t}_{\mathbf{O P}}(t)$ if $Z_{M_{t}}=A_{M_{t}}$ or if $Z_{M_{t}} \neq A_{M_{t}}$ but $F_{M_{t}} \neq A_{M_{t}}$. If $F_{M_{t}}=A_{M_{t}}$ there is no possible dimension reduction.

## References

1. Deville JC, Särndal CE. Calibration estimators in survey sampling. J Amer Statist Assoc. 1992;87(418):376-382.
2. Rueda MM, Martínez S, Martínez H, Arcos A. Estimation of the distribution function with calibration methods. J Stat Plan Infer. 2007;137(2):435-448.
3. Martínez S, Rueda MM, Martínez H, Arcos A. Optimal dimension and optimal auxiliary vector to construct calibration estimators of the distribution function. J Comput Appl Math. 2017;318:444-459.
4. Sedransk N, Sedransk J. Distinguishing among distributions using data from complex sample designs. J Am Stat Assoc. 1979;74(368):754-760.
5. Dickens R, Manning A. Has the national minimum wage reduced UK wage inequality?. J Roy Stat Soc A Sta. 2004;167(4):613-626.
6. Nickell S. Poverty and worklessness in Britain. The Economic Journal. 2004;114(494):C1-C25.
7. Machin S, Manning A, Rahman L. Where the minimum wage bites hard: Introduction of minimum wages to a low wage sector. Journal of the European Economic Association. 2003;1(1):154-180.
8. Estevao VM, Särndal CE. Survey estimates by calibration on complex auxiliary information. Int Stat Rev. 2006;74(2):127147.
9. Singh S. Generalized calibration approach for estimating variance in survey sampling. Ann I Stat Math. 2001;53(2):404-417.
10. Devaud D, Tillé Y. Deville and Särndal's calibration: revisiting a 25 years old successful optimization problem. Test. 2019;28(4):1033-1065.
11. Rueda MM. Comments on: Deville and Särndal's calibration: revisiting a 25 years old successful optimization problem. Test. 2019;28(4):1077-1081.
12. Kovacevic M. Calibration estimation of cumulative distribution and quantile functions from survey data. In: :139-144; 1997.
13. Harms T, Duchesne P. On calibration estimation for quantiles. Surv Methodol. 2006;32(1):37-52.
14. Wu C. Optimal calibration estimators in survey sampling. Biometrika. 2003;90(4):937-951.
15. Chen J, Wu C. Estimation of distribution function and quantiles using the model-calibrated pseudo empirical likelihood method. Statistica Sinica. 2002;12(4):1223-1239.
16. Breidt FJ, Opsomer JD, Johnson AA, Ranalli MG. Semiparametric model-assisted estimation for natural resource surveys. Survey Methodology. 2007;33(1):35.
17. Mayor-Gallego JA, Moreno-Rebollo JL, Jiménez-Gamero MD. Estimation of the finite population distribution function using a global penalized calibration method. Asta-Adv Stat Anal. 2019;103(1):1-35.
18. Martínez S, Rueda MM, Arcos A, Martínez H. Optimum calibration points estimating distribution functions. J Comput Appl Math. 2010;233(9):2265-2277.
19. Martínez S, Rueda MM, Arcos A, Martínez H, Sánchez-Borrego I. Post-stratified calibration method for estimating quantiles. Comput Stat Data An. 2011;55(1):838-851.
20. Martínez S, Rueda MM, Martínez H, Arcos A. Determining P optimum calibration points to construct calibration estimators of the distribution function. J Comput Appl Math. 2015;275:281-293.
21. Martínez S, Rueda MM, Illescas MD. The optimization problem of quantile and poverty measures estimation based on calibration. J Comput Appl Math. 2022;405:113054.
22. Nascimento Silva PLD, Skinner CJ. Variable selection for regression estimation in finite populations. Survey Methodology. 1997;23(1):23-32.
23. Chauvet G, Goga C. Asymptotic efficiency of the calibration estimator in a highdimensional data setting. Journal of Statistical Planning and Inference. 2022;217:177-187.
24. Chambers RL, Clark RG. Adaptive calibration for prediction of finite population totals. Survey Methodology. 2008;34(2):163-172.
25. McConville KS, Breidt FJ, Lee TMC, Moisen GG. Model assisted survey regression estimation with the LASSO. J. Surv. Stat. Methodol. 2017;5(2):131-158.
26. Guggemos F, Tillé Y. Penalized calibration in survey sampling: Design-based estimation assisted by mixed models. Journal of Statistical Planning and Inference. 2010;140(11):3199-3212.
27. Cardot H, Goga C, Shehzad M. Calibration and Partial Calibration on Principal Components when the Number of Auxiliary Variables is Large. Statistica Sinica. 2017;27(1):243-260.
28. Rota BJ. Variance estimation in two-step calibration for nonresponse adjustment. South African Statistical Journal. 2017;51(2):361-374.
29. Singh HP, Singh S, Kozak MA. A family of estimators of finite-population distribution function using auxiliary information. Acta Appl Math. 2008;104(2):115-130.
30. Arcos A, Martínez S, Rueda MM, Martínez H. Distribution function estimates from dual frame context. J Comput Appl Math. 2017;318:242-252.
31. Rao JNK, Kovar JG, Mantel HJ. On estimating distribution functions and quantiles from survey data using auxiliary information. Biometrika. 1986;77(2):365-375.
32. Chambers RL, Dunstan R. Estimating distribution functions from survey data. Biometrika. 1986;73(3):597-604.

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